## CS70: Jean Walrand: Lecture 16.

Events, Conditional Probability, Independence, Bayes' Rule

- 1. Probability Basics Review
- Events
- 3. Conditional Probability
- 4. Independence of Events
- 5. Bayes' Rule

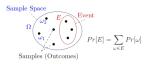
# Probability of exactly one 'heads' in two coin flips?

Idea: Sum the probabilities of all the different outcomes that have exactly one 'heads': HT, TH.

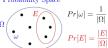
This leads to a definition!

## Definition:

- ▶ An **event**, E, is a subset of outcomes:  $E \subset \Omega$ .
- ▶ The **probability of** *E* is defined as  $Pr[E] = \sum_{\omega \in E} Pr[\omega]$ .



Uniform Probability Space



## **Probability Basics Review**

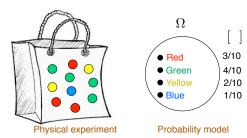
## Setup:

- Random Experiment. Flip a fair coin twice.
- Probability Space.
  - Sample Space: Set of outcomes, Ω.
    Ω = {HH, HT, TH, TT}

(Note: Not  $\Omega = \{H, T\}$  with two picks!)

- ► **Probability:**  $Pr[\omega]$  for all  $\omega \in \Omega$ .  $Pr[HH] = \cdots = Pr[TT] = 1/4$ 
  - 1.  $0 \le Pr[\omega] \le 1$ .
  - 2.  $\sum_{\omega \in \Omega} Pr[\omega] = 1$ .

# Event: Example



$$\Omega = \{ \text{Red, Green, Yellow, Blue} \}$$

$$Pr[\text{Red}] = \frac{3}{10}, Pr[\text{Green}] = \frac{4}{10}, \text{ etc.}$$

$$E = \{\textit{Red}, \textit{Green}\} \Rightarrow \textit{Pr}[E] = \frac{3+4}{10} = \frac{3}{10} + \frac{4}{10} = \textit{Pr}[\mathsf{Red}] + \textit{Pr}[\mathsf{Green}].$$

# 

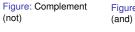


Figure: Intersection (and) Figure: Symmetric difference (only one)

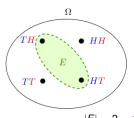
# Probability of exactly one heads in two coin flips?

Sample Space,  $\Omega = \{HH, HT, TH, TT\}.$ 

Uniform probability space:

 $Pr[HH] = Pr[HT] = Pr[TH] = Pr[TT] = \frac{1}{4}$ 

Event, *E*, "exactly one heads": {*TH*, *HT*}.



$$Pr[E] = \sum_{\omega \in E} Pr[\omega] = \frac{|E|}{|\Omega|} = \frac{2}{4} = \frac{1}{2}.$$

## Example: 20 coin tosses.

## 20 coin tosses

Sample space:  $\Omega=$  set of 20 fair coin tosses.  $\Omega=\{T,H\}^{20}\equiv\{0,1\}^{20};\ |\Omega|=2^{20}.$ 

- What is more likely?

  - $\omega_2 := (1,0,1,1,0,0,0,1,0,1,0,1,1,0,1,1,1,0,0,0)$ ?

Answer: Both are equally likely:  $Pr[\omega_1] = Pr[\omega_2] = \frac{1}{|\Omega|}$ .

- ▶ What is more likely?
- $(E_1)$  Twenty Hs out of twenty, or
- (E<sub>2</sub>) Ten Hs out of twenty?

Answer: Ten Hs out of twenty.

Why? There are many sequences of 20 tosses with ten Hs; only one with twenty Hs.  $\Rightarrow Pr[E_1] = \frac{1}{|\Omega|} \ll Pr[E_2] = \frac{|E_2|}{|\Omega|}$ .

$$|E_2| = {20 \choose 10} = 184,756.$$

# Exactly 50 heads in 100 coin tosses.

Sample space:  $\Omega = \text{set of } 100 \text{ coin tosses} = \{H, T\}^{100}.$  $|\Omega| = 2 \times 2 \times \cdots \times 2 = 2^{100}.$ 

Uniform probability space:  $Pr[\omega] = \frac{1}{2^{100}}$ .

Event  $\emph{E}=$  "100 coin tosses with exactly 50 heads"

|E|?

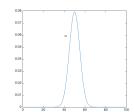
Choose 50 positions out of 100 to be heads.

 $|E| = \binom{100}{50}$ .

$$Pr[E] = \frac{\binom{100}{50}}{2^{100}}.$$

# Probability of *n* heads in 100 coin tosses.

$$\Omega = \{H, T\}^{100}; \ |\Omega| = 2^{100}.$$



Event  $E_n = 'n \text{ heads'}; |E_n| = \binom{100}{n}$ 

$$p_n := Pr[E_n] = \frac{|E_n|}{|\Omega|} = \frac{\binom{100}{n}}{2^{100}}$$

## Observe:

- ► Concentration around mean: Law of Large Numbers;
- Bell-shape: Central Limit Theorem.

## Calculation.

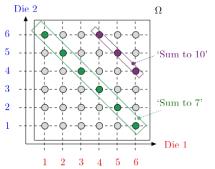
Stirling formula (for large *n*):

$$n! \approx \sqrt{2\pi n} \left(\frac{n}{e}\right)^n$$
.

$$\binom{2n}{n} \approx \frac{\sqrt{4\pi n} (2n/e)^{2n}}{[\sqrt{2\pi n} (n/e)^n]^2} \approx \frac{4^n}{\sqrt{\pi n}}.$$

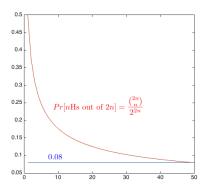
$$Pr[E] = \frac{|E|}{|\Omega|} = \frac{|E|}{2^{2n}} = \frac{1}{\sqrt{\pi n}} = \frac{1}{\sqrt{50\pi}} \approx .08.$$

## Roll a red and a blue die.



$$Pr[\text{Sum to 7}] = \frac{6}{36}$$
  $Pr[\text{Sum to 10}] = \frac{3}{36}$ 

# Exactly 50 heads in 100 coin tosses.



## Probability is Additive

## **Theorem**

(a) If events A and B are disjoint, i.e.,  $A \cap B = \emptyset$ , then

$$Pr[A \cup B] = Pr[A] + Pr[B].$$

(b) If events  $A_1, ..., A_n$  are pairwise disjoint, i.e.,  $A_k \cap A_m = \emptyset, \forall k \neq m$ , then

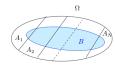
$$Pr[A_1 \cup \cdots \cup A_n] = Pr[A_1] + \cdots + Pr[A_n].$$

### Proof:

Obvious.

# Total probability

Assume that  $\Omega$  is the union of the disjoint sets  $A_1, \ldots, A_N$ .



Then,

$$Pr[B] = Pr[A_1 \cap B] + \cdots + Pr[A_N \cap B].$$

Indeed, *B* is the union of the disjoint sets  $A_n \cap B$  for n = 1, ..., N.

## Consequences of Additivity

## Theorem

(a) 
$$Pr[A \cup B] = Pr[A] + Pr[B] - Pr[A \cap B];$$
 (inclusion-exclusion property)

(b) 
$$Pr[A_1 \cup \cdots \cup A_n] \leq Pr[A_1] + \cdots + Pr[A_n];$$

(union bound)

(c) If  $A_1,\dots A_N$  are a partition of  $\Omega$ , i.e., pairwise disjoint and  $\cup_{m=1}^N A_m = \Omega$ , then

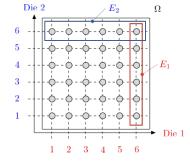
$$Pr[B] = Pr[B \cap A_1] + \cdots + Pr[B \cap A_N].$$
 (law of total probability)

## Proof:

(b) is obvious.

See next two slides for (a) and (c).

## Roll a Red and a Blue Die.



$$|E_1 \cup E_2| = |E_1| + |E_2| - |E_1 \cap E_2|$$

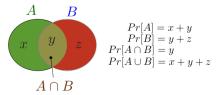
 $E_1$  = 'Red die shows 6';  $E_2$  = 'Blue die shows 6'

 $E_1 \cup E_2 =$  'At least one die shows 6'

$$Pr[E_1] = \frac{6}{36}, Pr[E_2] = \frac{6}{36}, Pr[E_1 \cup E_2] = \frac{11}{36}.$$

## Inclusion/Exclusion

$$Pr[A \cup B] = Pr[A] + Pr[B] - Pr[A \cap B]$$



# Conditional probability: example.

Two coin flips. First flip is heads. Probability of two heads?  $\Omega = \{HH, HT, TH, TT\}$ ; Uniform probability space. Event A =first flip is heads:  $A = \{HH, HT\}$ .



New sample space: A; uniform still.



Event B = two heads.

The probability of two heads if the first flip is heads.

The probability of B given A is 1/2.

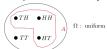
## A similar example.

Two coin flips. At least one of the flips is heads.

 $\rightarrow \text{Probability of two heads?}$ 

 $\Omega = \{HH, HT, TH, TT\}$ ; uniform.

Event A = at least one flip is heads.  $A = \{HH, HT, TH\}$ .



New sample space: A; uniform still.

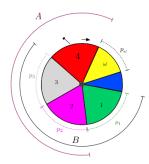


Event B = two heads.

The probability of two heads if at least one flip is heads. The probability of B given A is 1/3.

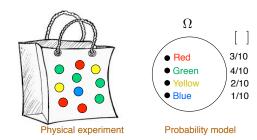
# Yet another non-uniform example

Consider  $\Omega = \{1, 2, ..., N\}$  with  $Pr[n] = p_n$ . Let  $A = \{2, 3, 4\}, B = \{1, 2, 3\}$ .



$$Pr[A|B] = \frac{p_2 + p_3}{p_1 + p_2 + p_3} = \frac{Pr[A \cap B]}{Pr[B]}$$

## Conditional Probability: A non-uniform example



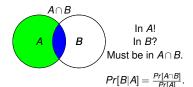
 $\Omega = \{ \text{Red, Green, Yellow, Blue} \}$ 

$$Pr[\text{Red} | \text{Red or Green}] = \frac{3}{7} = \frac{Pr[\text{Red} \cap (\text{Red or Green})]}{Pr[\text{Red or Green}]}$$

# Conditional Probability.

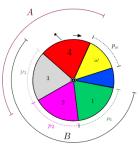
**Definition:** The **conditional probability** of *B* given *A* is

$$Pr[B|A] = \frac{Pr[A \cap B]}{Pr[A]}$$



## Another non-uniform example

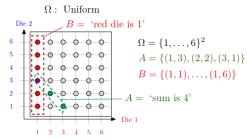
Consider  $\Omega = \{1, 2, ..., N\}$  with  $Pr[n] = p_n$ . Let  $A = \{3, 4\}, B = \{1, 2, 3\}$ .



$$Pr[A|B] = \frac{p_3}{p_1 + p_2 + p_3} = \frac{Pr[A \cap B]}{Pr[B]}$$

# More fun with conditional probability.

Toss a red and a blue die, sum is 4, What is probability that red is 1?



$$Pr[B|A] = \frac{|B \cap A|}{|A|} = \frac{1}{3}$$
; versus  $Pr[B] = 1/6$ .

B is more likely given A.

## Yet more fun with conditional probability.

Toss a red and a blue die, sum is 7, what is probability that red is 1?

# $\Omega: \text{ Uniform}$ $Die 2 \qquad B = \text{`red die is 1'}$ $6 \qquad \Omega = \{1, \dots, 6\}^2$ $A = \{(1, 6), \dots, (6,1)\}$ $B = \{(1, 1), \dots, (1, 6)\}$ A = `sum is 7' $1 \qquad Die 1$

$$Pr[B|A] = \frac{|B \cap A|}{|A|} = \frac{1}{6}$$
; versus  $Pr[B] = \frac{1}{6}$ .

Observing A does not change your mind about the likelihood of B.

## **Product Rule**

Recall the definition:

$$Pr[B|A] = \frac{Pr[A \cap B]}{Pr[A]}.$$

Hence,

$$Pr[A \cap B] = Pr[A]Pr[B|A].$$

Consequently,

$$Pr[A \cap B \cap C] = Pr[(A \cap B) \cap C]$$

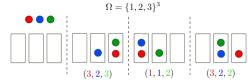
$$= Pr[A \cap B]Pr[C|A \cap B]$$

$$= Pr[A]Pr[B|A]Pr[C|A \cap B].$$

## Emptiness..

Suppose I toss 3 balls into 3 bins.

A = ``1st bin empty''; B = ``2nd bin empty.'' What is <math>Pr[A|B]?



 $\omega = (\text{bin of red ball}, \text{bin of blue ball}, \text{bin of green ball})$ 

$$Pr[B] = Pr[\{(a,b,c) \mid a,b,c \in \{1,3\}] = Pr[\{1,3\}^3] = \frac{8}{27}$$

$$Pr[A \cap B] = Pr[(3,3,3)] = \frac{1}{27}$$

$$Pr[A|B] = \frac{Pr[A \cap B]}{Pr[B]} = \frac{(1/27)}{(8/27)} = 1/8$$
; vs.  $Pr[A] = \frac{8}{27}$ .

 $\emph{A}$  is less likely given  $\emph{B}$ : If second bin is empty the first is more likely to have balls in it.

## **Product Rule**

**Theorem** Product Rule

Let  $A_1, A_2, \dots, A_n$  be events. Then

$$Pr[A_1 \cap \cdots \cap A_n] = Pr[A_1]Pr[A_2|A_1] \cdots Pr[A_n|A_1 \cap \cdots \cap A_{n-1}].$$

**Proof:** By induction.

Assume the result is true for n. (It holds for n = 2.) Then,

$$\begin{aligned} & Pr[A_{1} \cap \dots \cap A_{n} \cap A_{n+1}] \\ & = Pr[A_{1} \cap \dots \cap A_{n}] Pr[A_{n+1} | A_{1} \cap \dots \cap A_{n}] \\ & = Pr[A_{1}] Pr[A_{2} | A_{1}] \dots Pr[A_{n} | A_{1} \cap \dots \cap A_{n-1}] Pr[A_{n+1} | A_{1} \cap \dots \cap A_{n}], \end{aligned}$$

so that the result holds for n+1.

## Gambler's fallacy.

Flip a fair coin 51 times.

A = "first 50 flips are heads"

B = "the 51st is heads"

Pr[B|A] ?

 $\textit{A} = \{\textit{HH} \cdots \textit{HT}, \textit{HH} \cdots \textit{HH}\}$ 

 $\textit{B} \cap \textit{A} = \{\textit{HH} \cdots \textit{HH}\}$ 

Uniform probability space.

$$Pr[B|A] = \frac{|B \cap A|}{|A|} = \frac{1}{2}.$$

Same as Pr[B].

The likelihood of 51st heads does not depend on the previous flips.

## Correlation

An example.

Random experiment: Pick a person at random.

Event A: the person has lung cancer.

Event B: the person is a heavy smoker.

Fact:

 $\Box$ 

$$Pr[A|B] = 1.17 \times Pr[A].$$

Conclusion:

- Smoking increases the probability of lung cancer by 17%.
- Smoking causes lung cancer.

## Correlation

Event *A*: the person has lung cancer. Event *B*: the person is a heavy smoker.  $Pr[A|B] = 1.17 \times Pr[A]$ .

A second look.

Note that

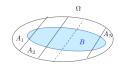
$$Pr[A|B] = 1.17 \times Pr[A] \Leftrightarrow \frac{Pr[A \cap B]}{Pr[B]} = 1.17 \times Pr[A]$$
$$\Leftrightarrow Pr[A \cap B] = 1.17 \times Pr[A]Pr[B]$$
$$\Leftrightarrow Pr[B|A] = 1.17 \times Pr[B].$$

## Conclusion:

- ▶ Lung cancer increases the probability of smoking by 17%.
- ► Lung cancer causes smoking. Really?

# Total probability

Assume that  $\Omega$  is the union of the disjoint sets  $A_1, \ldots, A_N$ .



Then,

$$Pr[B] = Pr[A_1 \cap B] + \cdots + Pr[A_N \cap B].$$

Indeed, B is the union of the disjoint sets  $A_n \cap B$  for  $n = 1, \dots, N$ . Thus,

$$Pr[B] = Pr[A_1]Pr[B|A_1] + \cdots + Pr[A_N]Pr[B|A_N].$$

## Causality vs. Correlation

Events A and B are positively correlated if

$$Pr[A \cap B] > Pr[A]Pr[B].$$

(E.g., smoking and lung cancer.)

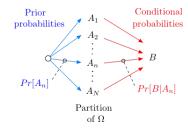
A and B being positively correlated does not mean that A causes B or that B causes A.

## Other examples:

- ► Tesla owners are more likely to be rich. That does not mean that poor people should buy a Tesla to get rich.
- People who go to the opera are more likely to have a good career. That does not mean that going to the opera will improve your career.
- Rabbits eat more carrots and do not wear glasses. Are carrots good for eyesight?

# Total probability

Assume that  $\Omega$  is the union of the disjoint sets  $A_1, \ldots, A_N$ .



$$Pr[B] = Pr[A_1]Pr[B|A_1] + \cdots + Pr[A_N]Pr[B|A_N].$$

## **Proving Causality**

Proving causality is generally difficult. One has to eliminate external causes of correlation and be able to test the cause/effect relationship (e.g., randomized clinical trials).

## Some difficulties:

- ► A and B may be positively correlated because they have a common cause. (E.g., being a rabbit.)
- If B precedes A, then B is more likely to be the cause. (E.g., smoking.) However, they could have a common cause that induces B before A. (E.g., smart, CS70, Tesla.)

More about such questions later. For fun, check "N. Taleb: Fooled by randomness."

## Is you coin loaded?

Your coin is fair w.p. 1/2 or such that Pr[H] = 0.6, otherwise.

You flip your coin and it yields heads.

What is the probability that it is fair?

## Analysis:

A = 'coin is fair', B = 'outcome is heads'

We want to calculate P[A|B].

We know P[B|A] = 1/2,  $P[B|\bar{A}] = 0.6$ ,  $Pr[A] = 1/2 = Pr[\bar{A}]$ Now.

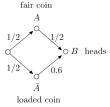
$$Pr[B] = Pr[A \cap B] + Pr[\bar{A} \cap B] = Pr[A]Pr[B|A] + Pr[\bar{A}]Pr[B|\bar{A}]$$
  
=  $(1/2)(1/2) + (1/2)0.6 = 0.55$ .

Thus,

$$Pr[A|B] = \frac{Pr[A]Pr[B|A]}{Pr[B]} = \frac{(1/2)(1/2)}{(1/2)(1/2) + (1/2)0.6} \approx 0.45.$$

## Is you coin loaded?

A picture:



Imagine 100 situations, among which

m := 100(1/2)(1/2) are such that  $\overline{A}$  and  $\overline{B}$  occur and n := 100(1/2)(0.6) are such that  $\overline{A}$  and  $\overline{B}$  occur.

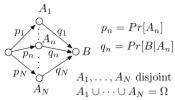
Thus, among the m+n situations where B occurred, there are m where A occurred.

Hence,

$$Pr[A|B] = \frac{m}{m+n} = \frac{(1/2)(1/2)}{(1/2)(1/2) + (1/2)0.6}.$$

## Bayes Rule

Another picture: We imagine that there are N possible causes  $A_1, \ldots, A_N$ .



Imagine 100 situations, among which  $100p_nq_n$  are such that  $A_n$  and B occur, for n = 1, ..., N.

Thus, among the  $100 \sum_{m} p_{m} q_{m}$  situations where *B* occurred, there are  $100 p_{n} q_{n}$  where  $A_{n}$  occurred.

Hence,

$$Pr[A_n|B] = \frac{p_n q_n}{\sum_m p_m q_m}$$

## Independence

**Definition:** Two events A and B are **independent** if

$$Pr[A \cap B] = Pr[A]Pr[B].$$

## Examples:

- ▶ When rolling two dice, A = sum is 7 and B = red die is 1 are independent;
- ► When rolling two dice, *A* = sum is 3 and *B* = red die is 1 are not independent:
- ► When flipping coins, *A* = coin 1 yields heads and *B* = coin 2 yields tails are independent;
- ► When throwing 3 balls into 3 bins, *A* = bin 1 is empty and *B* = bin 2 is empty are not independent;

# Why do you have a fever?



$$\textit{Pr}[\text{Flu}|\text{High Fever}] = \frac{0.15 \times 0.80}{0.15 \times 0.80 + 10^{-8} \times 1 + 0.85 \times 0.1} \approx 0.58$$

$$\textit{Pr}[\text{Ebola}|\text{High Fever}] = \frac{10^{-8} \times 1}{0.15 \times 0.80 + 10^{-8} \times 1 + 0.85 \times 0.1} \approx 5 \times 10^{-8}$$

$$Pr[\text{Other}|\text{High Fever}] = \frac{0.85 \times 0.1}{0.15 \times 0.80 + 10^{-8} \times 1 + 0.85 \times 0.1} \approx 0.42$$

These are the posterior probabilities. One says that 'Flu' is the Most Likely a Posteriori (MAP) cause of the high fever.

## Independence and conditional probability

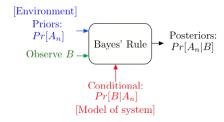
Fact: Two events A and B are independent if and only if

$$Pr[A|B] = Pr[A].$$

Indeed: 
$$Pr[A|B] = \frac{Pr[A \cap B]}{Pr[B]}$$
, so that

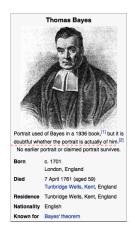
$$Pr[A|B] = Pr[A] \Leftrightarrow \frac{Pr[A \cap B]}{Pr[B]} = Pr[A] \Leftrightarrow Pr[A \cap B] = Pr[A]Pr[B].$$

# Bayes' Rule Operations



Bayes' Rule is the canonical example of how information changes our opinions.

# **Thomas Bayes**



Source: Wikipedia.

# Bayes Rule.



Using Bayes' rule, we find

$$P[A|B] = \frac{0.0016 \times 0.80}{0.0016 \times 0.80 + 0.9984 \times 0.10} = .013.$$

A 1.3% chance of prostate cancer with a positive PSA test.

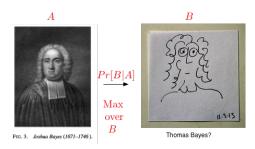
Surgery anyone?

Impotence...

Incontinence..

Death.

# **Thomas Bayes**



A Bayesian picture of Thomas Bayes.

## Summary

Events, Conditional Probability, Independence, Bayes' Rule

Key Ideas:

► Conditional Probability:

$$Pr[A|B] = \frac{Pr[A \cap B]}{Pr[B]}$$

- ▶ Independence:  $Pr[A \cap B] = Pr[A]Pr[B]$ .
- ▶ Bayes' Rule:

$$Pr[A_n|B] = \frac{Pr[A_n]Pr[B|A_n]}{\sum_m Pr[A_m]Pr[B|A_m]}.$$

 $Pr[A_n|B] = posterior probability; Pr[A_n] = prior probability$ .

► All these are possible:

$$Pr[A|B] < Pr[A]; Pr[A|B] > Pr[A]; Pr[A|B] = Pr[A].$$

## Testing for disease.

Let's watch TV!!

Random Experiment: Pick a random male.

Outcomes: (test, disease)

A - prostate cancer.

B - positive PSA test.

- ightharpoonup Pr[A] = 0.0016, (.16 % of the male population is affected.)
- ▶ Pr[B|A] = 0.80 (80% chance of positive test with disease.)
- ▶  $Pr[B|\overline{A}] = 0.10$  (10% chance of positive test without disease.)

From http://www.cpcn.org/01\_psa\_tests.htm and http://seer.cancer.gov/statfacts/html/prost.html (10/12/2011.)

Positive PSA test (B). Do I have disease?

Pr[A|B]???