

CS70: Jean Walrand: Lecture 23.

Conditional Expectation

1. Review: LR and LLSE
2. Conditional expectation
3. Applications: Diluting, Mixing, Rumors
4. CE = MMSE

Review: LLSE and LR

Definitions Let X and Y be RVs on Ω .

- ▶ **Covariance:** $\text{cov}(X, Y) := E[XY] - E[X]E[Y]$
- ▶ **LLSE:** $L[Y|X] = a + bX$ where a, b minimize $E[(Y - a - bX)^2]$.

We saw that

$$L[Y|X] = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}[X]}(X - E[X]).$$

Then,

$$E[(Y - L[Y|X])^2] = \text{var}(Y) - \text{cov}(X, Y)^2 / \text{var}(X).$$

Non-Bayesian (LR): We are given samples $(X_1, Y_1), \dots, (X_K, Y_K)$, no distribution.

We define the RVs (X, Y) so that

$$\Pr[(X, Y) = (X_k, Y_k)] = 1/K, k = 1, \dots, K.$$

Then, as before.

Review: LLSE and LR

Consider the non-Bayesian case: sample $(X_1, Y_1), \dots, (X_K, Y_K)$.

Then,

$$L[Y|X] = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]).$$

Here,

$$E[X] = \frac{1}{K} \sum_{k=1}^K X_k$$

$$E[Y] = \frac{1}{K} \sum_{k=1}^K Y_k$$

$$E[X^2] = \frac{1}{K} \sum_{k=1}^K X_k^2$$

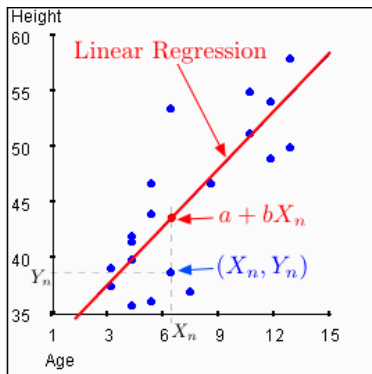
$$E[XY] = \frac{1}{K} \sum_{k=1}^K X_k Y_k$$

$$\text{cov}(X, Y) = E[XY] - E[X]E[Y]$$

$$\text{var}(X) = E[X^2] - E[X]^2.$$

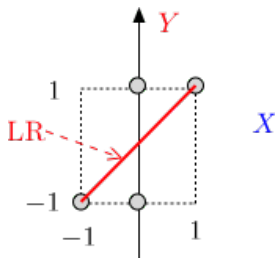
Linear Regression Examples

Example 1:



Linear Regression Examples

Example 2: Four equally likely values of (X, Y) , or four samples.



We find:

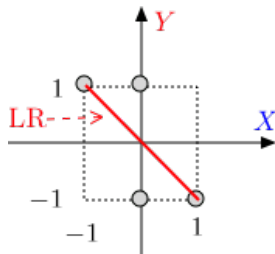
$$E[X] = 0; E[Y] = 0; E[X^2] = 1/2; E[XY] = 1/2;$$

$$\text{var}[X] = E[X^2] - E[X]^2 = 1/2; \text{cov}(X, Y) = E[XY] - E[X]E[Y] = 1/2;$$

$$\text{LR: } \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}[X]}(X - E[X]) = X.$$

Linear Regression Examples

Example 3: Four equally likely values of (X, Y) , or four samples.



We find:

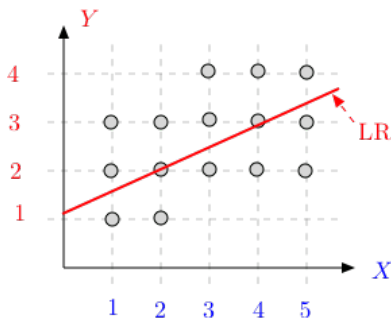
$$E[X] = 0; E[Y] = 0; E[X^2] = 1/2; E[XY] = -1/2;$$

$$\text{var}[X] = E[X^2] - E[X]^2 = 1/2; \text{cov}(X, Y) = E[XY] - E[X]E[Y] = -1/2;$$

$$\text{LR: } \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}[X]}(X - E[X]) = -X.$$

Linear Regression Examples

Example 4: Equally likely values of (X, Y) , or samples.



We find:

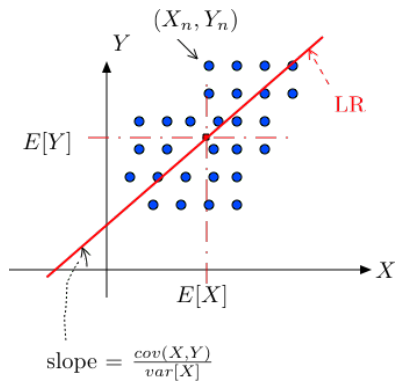
$$E[X] = 3; E[Y] = 2.5; E[X^2] = (3/15)(1 + 2^2 + 3^2 + 4^2 + 5^2) = 11;$$

$$E[XY] = (1/15)(1 \times 1 + 1 \times 2 + \dots + 5 \times 4) = 8.4;$$

$$\text{var}[X] = 11 - 9 = 2; \text{cov}(X, Y) = 8.4 - 3 \times 2.5 = 0.9;$$

$$\text{LR: } \hat{Y} = 2.5 + \frac{0.9}{2}(X - 3) = 1.15 + 0.45X.$$

LR: Another Figure



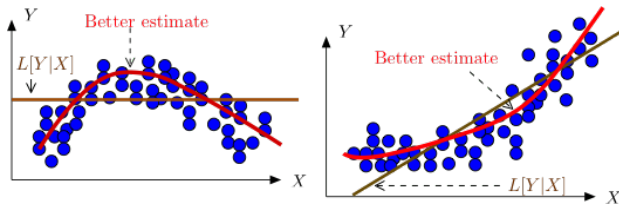
Note that

- ▶ the LR line goes through $(E[X], E[Y])$
- ▶ its slope is $\frac{\text{cov}(X, Y)}{\text{var}(X)}$.

Conditional Expectation: Motivation

There are many situations where a good guess about Y given X is not linear.

E.g., (diameter of object, weight), (school years, income), (PSA level, cancer risk).

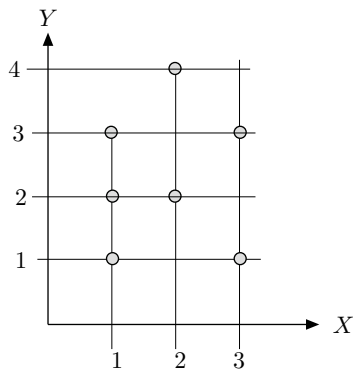


Our goal: Derive the best estimate of Y given X !

That is, find the function $g(\cdot)$ so that $g(X)$ is the best guess about Y given X .

Ambitious! Can it be done? Amazingly, yes!

Conditional Expectation: Intuition

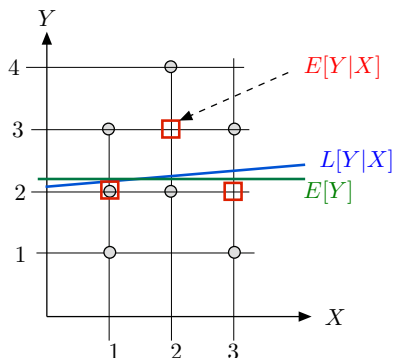


Without any observation, our guess for Y is $E[Y] = 2.3$.

Assume now we observe X . We can calculate $L[Y|X] = a + bX \approx 2.1 + 0.1x$.

A better guess when $X = 1$ is 2; when $X = 2$: 3; when $X = 3$: 2.

Conditional Expectation: Intuition



Here, $E[Y|X = 1]$ is the mean value of Y given that $X = 1$. Also, $E[Y|X = 2]$ is the mean value of Y given that $X = 2$ and $E[Y|X = 3]$ is the mean value of Y given that $X = 3$.

When we know that $X = 1$, Y has a new distribution: Y is uniform in $\{1, 2, 3\}$.

Thus, our guess is $E[Y|X = 1] = 1(1/3) + 2(1/3) + 3(1/3) = 2$.

Conditional Expectation

Definition Let X and Y be RVs on Ω . The **conditional expectation** of Y given X is defined as

$$E[Y|X] = g(X)$$

where

$$g(x) := E[Y|X = x] := \sum_y y Pr[Y = y|X = x],$$

with $Pr[Y = y|X = x] := \frac{Pr[X=x, Y=y]}{Pr[X=x]}$.

Theorem: $E[Y|X]$ is the best guess about Y given X .

That is, for any function $h(\cdot)$, one has

$$E[(Y - h(X))^2] \geq E[(Y - E[Y|X])^2].$$

Proof: Later.

Calculating $E[Y|X]$

Let X, Y, Z be i.i.d. with mean 0 and variance 1. We want to calculate

$$E[2 + 5X + 7XY + 11X^2 + 13X^3Z^2|X].$$

We find

$$\begin{aligned} E[2 + 5X + 7XY + 11X^2 + 13X^3Z^2|X] &= 2 + 5X + 7XE[Y|X] + 11X^2 + 13X^3E[Z^2|X] \\ &= 2 + 5X + 7XE[Y] + 11X^2 + 13X^3E[Z^2] \\ &= 2 + 5X + 11X^2 + 13X^3(\text{var}[Z] + E[Z]^2) \\ &= 2 + 5X + 11X^2 + 13X^3. \end{aligned}$$

Projection Property

The claim is that

$$E[(Y - E[Y|X])f(X)] = 0, \forall f(\cdot).$$

That is,

$$E[Yf(X)] = E[E[Y|X]f(X)]$$

In particular, choosing $f(x) = 1$, we get

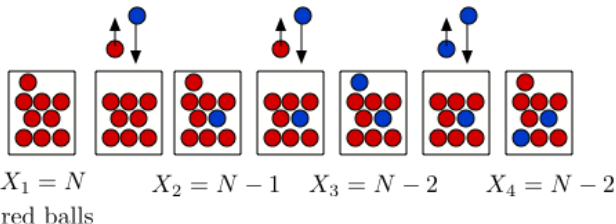
$$E[Y] = E[E[Y|X]].$$

Proof:

$$\begin{aligned} E[E[Y|X]f(X)] &= \sum_x E[Y|X = x]f(x)Pr[X = x] \\ &= \sum_x [\sum_y yf(x)Pr[Y = y|X = x]]Pr[X = x] \\ &= \sum_x \sum_y yf(x)Pr[X = x, Y = y] \\ &= E[Yf(X)]. \end{aligned}$$



Application: Diluting



At each step, pick a ball from a well-mixed urn. Replace it with a blue ball. Let X_n be the number of red balls in the urn at step n . What is $E[X_n]$?

Given $X_n = m$, $X_{n+1} = m - 1$ w.p. m/N (if you pick a red ball) and $X_{n+1} = m$ otherwise. Hence,

$$E[X_{n+1} | X_n = m] = m - (m/N) = m(N - 1)/N = X_n \rho,$$

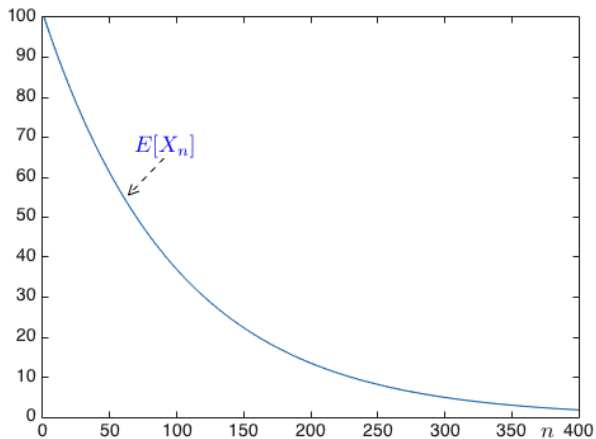
with $\rho := (N - 1)/N$. Consequently,

$$E[X_{n+1}] = E[E[X_{n+1} | X_n]] = \rho E[X_n], n \geq 1.$$

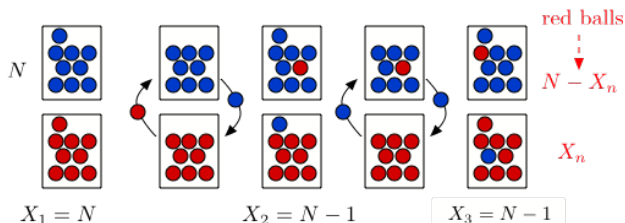
$$\implies E[X_n] = \rho^{n-1} E[X_1] = N \left(\frac{N-1}{N} \right)^{n-1}, n \geq 1.$$

Diluting

Here is a plot:



Application: Mixing



At each step, pick a ball from each well-mixed urn. We transfer them to the other urn. Let X_n be the number of red balls in the bottom urn at step n . What is $E[X_n]$?

Given $X_n = m$, $X_{n+1} = m + 1$ w.p. p and $X_{n+1} = m - 1$ w.p. q

where $p = (1 - m/N)^2$ (B goes up, R down) and $q = (m/N)^2$ (R goes up, B down).

Thus,

$$E[X_{n+1}|X_n] = X_n + p - q = X_n + 1 - 2X_n/N = 1 + \rho X_n, \quad \rho := (1 - 2/N).$$

Mixing

We saw that $E[X_{n+1}|X_n] = 1 + \rho X_n$, $\rho := (1 - 2/N)$. Hence,

$$E[X_{n+1}] = 1 + \rho E[X_n]$$

$$E[X_2] = 1 + \rho N; E[X_3] = 1 + \rho(1 + \rho N) = 1 + \rho + \rho^2 N$$

$$E[X_4] = 1 + \rho(1 + \rho + \rho^2 N) = 1 + \rho + \rho^2 + \rho^3 N$$

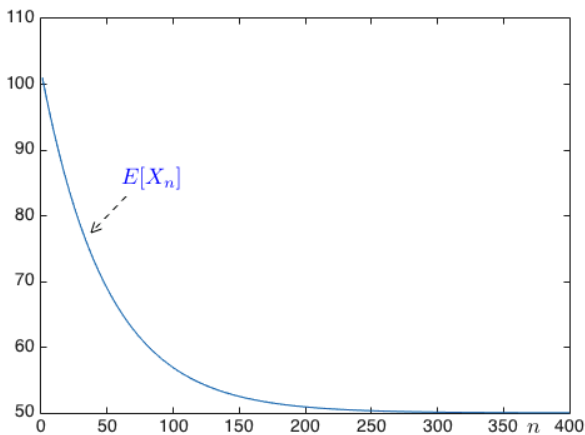
$$E[X_n] = 1 + \rho + \dots + \rho^{n-2} + \rho^{n-1} N.$$

Hence,

$$E[X_n] = \frac{1 - \rho^{n-1}}{1 - \rho} + \rho^{n-1} N, n \geq 1.$$

Application: Mixing

Here is the plot.



Application: Going Viral

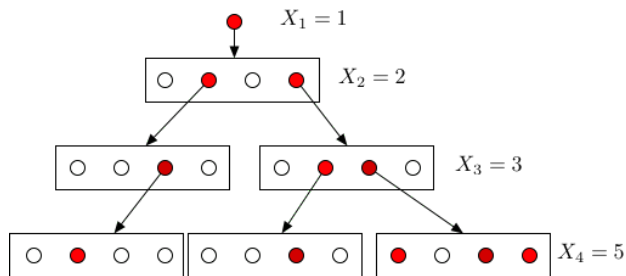
Consider a social network (e.g., Twitter).

You start a rumor (e.g., Walrand is really weird).

You have d friends. Each of your friend retweets w.p. p .

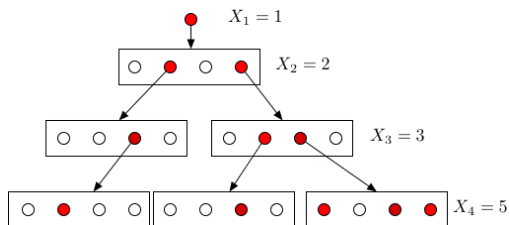
Each of your friends has d friends, etc.

Does the rumor spread? Does it die out (mercifully)?



In this example, $d = 4$.

Application: Going Viral



Fact: Let $X = \sum_{n=1}^{\infty} X_n$. Then, $E[X] < \infty$ iff $pd < 1$.

Proof:

Given $X_n = k$, $X_{n+1} = B(kd, p)$. Hence, $E[X_{n+1} | X_n = k] = kpd$.

Thus, $E[X_{n+1} | X_n] = pdX_n$. Consequently, $E[X_n] = (pd)^{n-1}$, $n \geq 1$.

If $pd < 1$, then $E[X_1 + \dots + X_n] \leq (1 - pd)^{-1} \implies E[X] \leq (1 - pd)^{-1}$.

If $pd \geq 1$, then for all C one can find n s.t.

$$E[X] \geq E[X_1 + \dots + X_n] \geq C.$$

□

In fact, one can show that $pd \geq 1 \implies \Pr[X = \infty] > 0$.

Application: Wald's Identity

Theorem Wald's Identity

Assume that X_1, X_2, \dots and Z are independent, where

Z takes values in $\{0, 1, 2, \dots\}$

and $E[X_n] = \mu$ for all $n \geq 1$.

Then,

$$E[X_1 + \dots + X_Z] = \mu E[Z].$$

Proof:

$$E[X_1 + \dots + X_Z | Z = k] = \mu k.$$

$$\text{Thus, } E[X_1 + \dots + X_Z | Z] = \mu Z.$$

$$\text{Hence, } E[X_1 + \dots + X_Z] = E[\mu Z] = \mu E[Z].$$



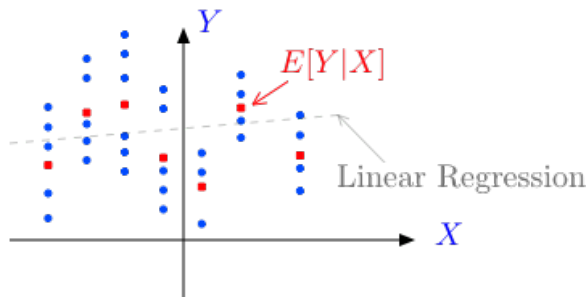
CE = MMSE

Theorem

$E[Y|X]$ is the 'best' guess about Y based on X .

Specifically, it is the function $g(X)$ of X that

minimizes $E[(Y - g(X))^2]$.



CE = MMSE

Theorem CE = MMSE

$g(X) := E[Y|X]$ is the function of X that minimizes $E[(Y - g(X))^2]$.

Proof:

Let $h(X)$ be any function of X . Then

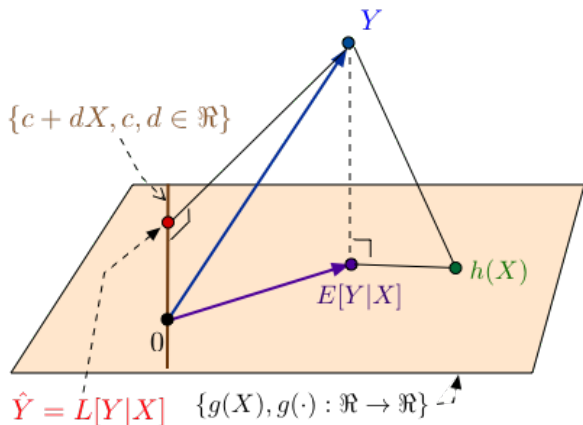
$$\begin{aligned} E[(Y - h(X))^2] &= E[(Y - g(X) + g(X) - h(X))^2] \\ &= E[(Y - g(X))^2] + E[(g(X) - h(X))^2] \\ &\quad + 2E[(Y - g(X))(g(X) - h(X))]. \end{aligned}$$

But,

$$E[(Y - g(X))(g(X) - h(X))] = 0 \text{ by the projection property.}$$

Thus, $E[(Y - h(X))^2] \geq E[(Y - g(X))^2]$. □

$E[Y|X]$ and $L[Y|X]$ as projections



$L[Y|X]$ is the projection of Y on $\{a + bX, a, b \in \mathbb{R}\}$: LLSE

$E[Y|X]$ is the projection of Y on $\{g(X), g(\cdot) : \mathbb{R} \rightarrow \mathbb{R}\}$: MMSE.

Summary

Conditional Expectation

- ▶ Definition: $E[Y|X] := \sum_y yPr[Y = y|X = x]$
- ▶ Properties: Linearity,
 $Y - E[Y|X] \perp h(X)$; $E[E[Y|X]] = E[Y]$
- ▶ Some Applications:
 - ▶ Calculating $E[Y|X]$
 - ▶ Diluting
 - ▶ Mixing
 - ▶ Rumors
 - ▶ Wald
- ▶ MMSE: $E[Y|X]$ minimizes $E[(Y - g(X))^2]$ over all $g(\cdot)$