CS70: Jean Walrand: Lecture 26.

Continuous Probability

- 1. Examples
- 2. Events
- 3. Continuous Random Variables
- 4. Expectation
- 5. Bayes' Rule
- 6. Multiple Random Variables

Continuous Probability - Pick a real number.

Choose a real number X, uniformly at random in [0, 1000].

What is the probability that X is exactly equal to $100\pi = 314.1592625...$? Well, ..., 0.



Let [a, b] denote the **event** that the point X is in the interval [a, b].

$$Pr[[a,b]] = \frac{\text{length of } [a,b]}{\text{length of } [0,L]} = \frac{b-a}{L} = \frac{b-a}{1000}$$

Intervals like $[a, b] \subseteq \Omega = [0, L]$ are **events.** More generally, events in this space are unions of intervals. Example: the event *A* - "within 50 of 0 or 1000" is $A = [0, 50] \cup [950, 1000]$. Thus,

$$Pr[A] = Pr[[0, 50]] + Pr[[950, 10000]] = \frac{1}{10}.$$

Continuous Probability - Pick a random real number.



Note: A **radical** change in approach. For a finite probability space, $\Omega = \{1, 2, ..., N\}$, we started with $Pr[\omega] = p_{\omega}$. We then defined $Pr[A] = \sum_{\omega \in A} p_{\omega}$ for $A \subset \Omega$. We used the same approach for countable Ω .

For a continuous space, e.g., $\Omega = [0, L]$, we cannot start with $Pr[\omega]$, because this will typically be 0. Instead, we start with Pr[A] for some events *A*. Here, we started with *A* = interval, or union of intervals.

Thus, the probability is a function from events to [0,1]. Can any function make sense? No! At least, it should be additive!. In our example, $Pr[[0,50] \cup [950,1000]] = Pr[[0,50]] + Pr[[950,1000]]$.

Shooting..

A James Bond example. In Spectre, Mr. Hinx is chasing Bond who is in a Aston Martin DB10. Hinx shoots at the DB10 and hits it at a random spot. What is the chance Hinx hits the gas tank? Assume the gas tank is a one foot circle and the DB10 is an

expensive 4×5 rectangle.



 $\Omega = \{(x, y) : x \in [0, 4], y \in [0, 5]\}.$

The size of the event is $\pi(1)^2 = \pi$. The "size" of the sample space which is 4×5 . Since uniform, probability of event is $\frac{\pi}{20}$.

Continuous Random Variables: CDF

 $Pr[a < X \le b]$ instead of Pr[X = a]. For all *a* and *b*: specifies the behavior! Simpler: $P[X \le x]$ for all *x*.

Cumulative probability Distribution Function of X is

$$F_X(x) = Pr[X \leq x]^1$$

$$Pr[a < X \le b] = Pr[X \le b] - Pr[X \le a] = F_X(b) - F_X(a).$$

Idea: two events $X \le b$ and $X \le a$. Difference is the event $a < X \le b$. Indeed: $\{X \le b\} \setminus \{X \le a\} = \{X \le b\} \cap \{X > a\} = \{a < X \le b\}$.

¹The subscript X reminds us that this corresponds to the RV X.

Example: CDF

Example: Value of X in [0, L] with L = 1000.

$$F_X(x) = \Pr[X \le x] = \begin{cases} 0 & \text{for } x < 0\\ \frac{x}{1000} & \text{for } 0 \le x \le 1000\\ 1 & \text{for } x > 1000 \end{cases}$$

Probability that *X* is within 50 of center:

$$Pr[450 < X \le 550] = Pr[X \le 550] - Pr[X \le 450]$$
$$= \frac{550}{1000} - \frac{450}{1000}$$
$$= \frac{100}{1000} = \frac{1}{10}$$



Example: hitting random location on gas tank. Random location on circle.



Random Variable: *Y* distance from center. Probability within *y* of center:

$$Pr[Y \le y] = \frac{\text{area of small circle}}{\text{area of dartboard}}$$
$$= \frac{\pi y^2}{\pi} = y^2.$$

Hence,

$$F_{Y}(y) = \Pr[Y \le y] = \begin{cases} 0 & \text{for } y < 0\\ y^{2} & \text{for } 0 \le y \le 1\\ 1 & \text{for } y > 1 \end{cases}$$

Calculation of event with dartboard..

Probability between .5 and .6 of center? Recall CDF.

$$F_{Y}(y) = \Pr[Y \le y] = \begin{cases} 0 & \text{for } y < 0\\ y^{2} & \text{for } 0 \le y \le 1\\ 1 & \text{for } y > 1 \end{cases}$$

$$Pr[0.5 < Y \le 0.6] = Pr[Y \le 0.6] - Pr[Y \le 0.5]$$

= F_Y(0.6) - F_Y(0.5)
= .36 - .25
= .11

Density function.

Is the dart more like to be (near) .5 or .1? Probability of "Near x" is $Pr[x < X \le x + \delta]$. Goes to 0 as δ goes to zero. Try

$$\frac{\Pr[x < X \le x + \delta]}{\delta}$$

The limit as δ goes to zero.

$$\lim_{\delta \to 0} \frac{\Pr[x < X \le x + \delta]}{\delta} = \lim_{\delta \to 0} \frac{\Pr[X \le x + \delta] - \Pr[X \le x]}{\delta}$$
$$= \lim_{\delta \to 0} \frac{F_X(x + \delta) - F_X(x)}{\delta}$$
$$= \frac{d(F_X(x))}{dx}.$$

Density

Definition: (Density) A **probability density function** for a random variable *X* with cdf $F_X(x) = Pr[X \le x]$ is the function $f_X(x)$ where

$$F_X(x) = \int_{-\infty}^x f_X(u) du.$$

Thus,

$$\Pr[X \in (x, x+\delta]] = F_X(x+\delta) - F_X(x) \approx f_X(x)\delta.$$

Examples: Density.

Example: uniform over interval [0,1000]

$$f_X(x) = F'_X(x) = \begin{cases} 0 & \text{for } x < 0\\ \frac{1}{1000} & \text{for } 0 \le x \le 1000\\ 0 & \text{for } x > 1000 \end{cases}$$

Example: uniform over interval [0, L]

$$f_X(x) = F'_X(x) = \begin{cases} 0 & \text{for } x < 0\\ \frac{1}{L} & \text{for } 0 \le x \le L\\ 0 & \text{for } x > L \end{cases}$$

Examples: Density.

Example: "Dart" board. Recall that

$$F_{Y}(y) = Pr[Y \le y] = \begin{cases} 0 & \text{for } y < 0\\ y^{2} & \text{for } 0 \le y \le 1\\ 1 & \text{for } y > 1 \end{cases}$$
$$f_{Y}(y) = F_{Y}'(y) = \begin{cases} 0 & \text{for } y < 0\\ 2y & \text{for } 0 \le y \le 1\\ 0 & \text{for } y > 1 \end{cases}$$

The cumulative distribution function (cdf) and probability distribution function (pdf) give full information. Use whichever is convenient.

Target



U[*a*,*b*]



$Expo(\lambda)$

The exponential distribution with parameter $\lambda > 0$ is defined by $f_X(x) = \lambda e^{-\lambda x} \mathbb{1}\{x \ge 0\}$

$$F_X(x) = \begin{cases} 0, & \text{if } x < 0\\ 1 - e^{-\lambda x}, & \text{if } x \ge 0. \end{cases}$$



Note that $Pr[X > t] = e^{-\lambda t}$ for t > 0.

Random Variables

Continuous random variable X, specified by

1.
$$F_X(x) = Pr[X \le x]$$
 for all x .
Cumulative Distribution Function (cdf).
 $Pr[a < X \le b] = F_X(b) - F_X(a)$
1.1 $0 \le F_X(x) \le 1$ for all $x \in \Re$.
1.2 $F_X(x) \le F_X(y)$ if $x \le y$.

2. Or $f_X(x)$, where $F_X(x) = \int_{-\infty}^x f_X(u) du$ or $f_X(x) = \frac{d(F_X(x))}{dx}$. **Probability Density Function (pdf).** $Pr[a < X \le b] = \int_a^b f_X(x) dx = F_X(b) - F_X(a)$ 2.1 $f_X(x) \ge 0$ for all $x \in \mathfrak{R}$. 2.2 $\int_{-\infty}^{\infty} f_X(x) dx = 1$.

Recall that $Pr[X \in (x, x + \delta)] \approx f_X(x)\delta$. Think of X taking discrete values $n\delta$ for n = ..., -2, -1, 0, 1, 2, ... with $Pr[X = n\delta] = f_X(n\delta)\delta$.

A Picture



The pdf $f_X(x)$ is a nonnegative function that integrates to 1. The cdf $F_X(x)$ is the integral of f_X .

$$Pr[x < X < x + \delta] \approx f_X(x)\delta$$
$$Pr[X \le x] = F_x(x) = \int_{-\infty}^x f_X(u)du$$

Some Examples

1. *Expo* is memoryless. Let $X = Expo(\lambda)$. Then, for s, t > 0,

$$Pr[X > t + s \mid X > s] = \frac{Pr[X > t + s]}{Pr[X > s]}$$
$$= \frac{e^{-\lambda(t+s)}}{e^{-\lambda s}} = e^{-\lambda t}$$
$$= Pr[X > t].$$

'Used is a good as new.'

2. Scaling *Expo*. Let $X = Expo(\lambda)$ and Y = aX for some a > 0. Then

$$\begin{aligned} \Pr[Y > t] &= \Pr[aX > t] = \Pr[X > t/a] \\ &= e^{-\lambda(t/a)} = e^{-(\lambda/a)t} = \Pr[Z > t] \text{ for } Z = Expo(\lambda/a). \end{aligned}$$

Thus, $a \times Expo(\lambda) = Expo(\lambda/a)$.

Some More Examples

3. Scaling Uniform. Let X = U[0, 1] and Y = a + bX where b > 0. Then,

$$Pr[Y \in (y, y+\delta)] = Pr[a+bX \in (y, y+\delta)] = Pr[X \in (\frac{y-a}{b}, \frac{y+\delta-a}{b})]$$
$$= Pr[X \in (\frac{y-a}{b}, \frac{y-a}{b} + \frac{\delta}{b})] = \frac{1}{b}\delta, \text{ for } 0 < \frac{y-a}{b} < 1$$
$$= \frac{1}{b}\delta, \text{ for } a < y < a+b.$$

Thus, $f_Y(y) = \frac{1}{b}$ for a < y < a+b. Hence, Y = U[a, a+b]. **4. Scaling pdf.** Let $f_X(x)$ be the pdf of X and Y = a+bX where b > 0. Then

$$Pr[Y \in (y, y+\delta)] = Pr[a+bX \in (y, y+\delta)] = Pr[X \in (\frac{y-a}{b}, \frac{y+\delta-a}{b}]$$
$$= Pr[Pr[X \in (\frac{y-a}{b}, \frac{y-a}{b} + \frac{\delta}{b}] = f_X(\frac{y-a}{b})\frac{\delta}{b}.$$

Now, the left-hand side is $f_Y(y)\delta$. Hence,

$$f_Y(y)=\frac{1}{b}f_X(\frac{y-a}{b}).$$

Expectation

Definition The **expectation** of a random variable X with pdf f(x) is defined as E

$$\Xi[X] = \int_{-\infty}^{\infty} x f_X(x) dx.$$

Justification: Say $X = n\delta$ w.p. $f_X(n\delta)\delta$. Then,

$$E[X] = \sum_{n} (n\delta) Pr[X = n\delta] = \sum_{n} (n\delta) f_X(n\delta) \delta = \int_{-\infty}^{\infty} x f_X(x) dx.$$

Indeed, for any *g*, one has $\int g(x) dx \approx \sum_n g(n\delta)\delta$. Choose $q(x) = x f_X(x).$ $g(n\delta)\delta$



Expectation of function of RV

Definition The expectation of a function of a random variable is defined as

$$E[h(X)] = \int_{-\infty}^{\infty} h(x) f_X(x) dx.$$

Justification: Say $X = n\delta$ w.p. $f_X(n\delta)\delta$. Then,

$$E[h(X)] = \sum_{n} h(n\delta) Pr[X = n\delta] = \sum_{n} h(n\delta) f_X(n\delta) \delta = \int_{-\infty}^{\infty} h(x) f_X(x) dx.$$

Indeed, for any g, one has $\int g(x) dx \approx \sum_n g(n\delta)\delta$. Choose $g(x) = h(x)f_X(x)$.

Fact Expectation is linear. Proof: As in the discrete case.

Variance

Definition: The **variance** of a continuous random variable *X* is defined as

$$var[X] = E((X - E(X))^2) = E(X^2) - (E(X))^2$$

= $\int_{-\infty}^{\infty} x^2 f(x) dx - \left(\int_{-\infty}^{\infty} x f(x) dx\right)^2.$

- Key fact: The sum of many small independent RVs has a Gaussian distribution.
- This is the Central Limit Theorem. (See later.)
- Examples: Binomial and Poisson suitably scaled.
- This explains why the Gaussian distribution (the bell curve) shows up everywhere.

Summary

Continuous Probability

- 1. pdf: $Pr[X \in (x, x + \delta]] = f_X(x)\delta$.
- 2. CDF: $Pr[X \le x] = F_X(x) = \int_{-\infty}^x f_X(y) dy$.
- 3. U[a,b], $Expo(\lambda)$, target.
- 4. Expectation: $E[X] = \int_{-\infty}^{\infty} x f_X(x) dx$.
- 5. Expectation of function: $E[h(X)] = \int_{-\infty}^{\infty} h(x) f_X(x) dx$.
- 6. Variance: $var[X] = E[(X E[X])^2] = E[X^2] E[X]^2$.
- 7. Gaussian: $\mathcal{N}(\mu, \sigma^2)$: $f_X(x) = \dots$ "bell curve"