

## CS70: Jean Walrand: Lecture 18.

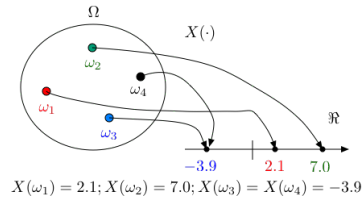
### Random Variables & Midterm 2 Probability Review

- ▶ Random Variables
- ▶ M2 Probability Review
- ▶ M2 Discrete Math Review: See Video (link given on Piazza)

### Random Variables.

A **random variable**,  $X$ , for an experiment with sample space  $\Omega$  is a **function**  $X : \Omega \rightarrow \mathbb{R}$ .

Thus,  $X(\cdot)$  assigns a real number  $X(\omega)$  to each  $\omega \in \Omega$ .



The function  $X(\cdot)$  is defined on the outcomes  $\Omega$ .

The function  $X(\cdot)$  is **not random, not a variable!**

What varies at random (from experiment to experiment)? The outcome!

### Random Variables

1. Random Variables.
2. Distributions.
3. Combining random variables.
4. Expectation

### Example 1 of Random Variable

Experiment: roll two dice.

Sample Space:  $\{(1,1), (1,2), \dots, (6,6)\} = \{1, \dots, 6\}^2$

Random Variable  $X$ : number of pips.

$$X(1,1) = 2$$

$$X(1,2) = 3,$$

$\vdots$

$$X(6,6) = 12,$$

$$X(a,b) = a + b, (a,b) \in \Omega.$$

### Questions about outcomes ...

Experiment: roll two dice.

Sample Space:  $\{(1,1), (1,2), \dots, (6,6)\} = \{1, \dots, 6\}^2$

How many pips?

Experiment: flip 100 coins.

Sample Space:  $\{HHH \dots H, THH \dots H, \dots, TTT \dots T\}$

How many heads in 100 coin tosses?

Experiment: choose a random student in cs70.

Sample Space:  $\{Adam, Jin, Bing, \dots, Angeline\}$

What midterm score?

Experiment: hand back assignments to 3 students at random.

Sample Space:  $\{123, 132, 213, 231, 312, 321\}$

How many students get back their own assignment?

In each scenario, each outcome gives a number.

The number is a (known) function of the outcome.

### Example 2 of Random Variable

Experiment: flip three coins

Sample Space:  $\{HHH, THH, HTH, TTH, HHT, THT, HTT, TTT\}$

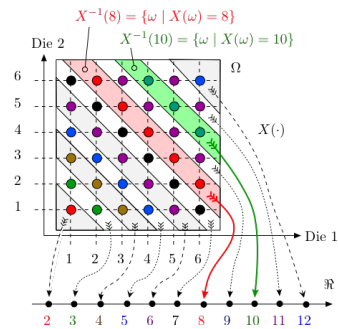
Winnings: if win 1 on heads, lose 1 on tails:  $X$

$$X(HHH) = 3 \quad X(THH) = 1 \quad X(HTH) = 1 \quad X(TTH) = -1$$

$$X(HHT) = 1 \quad X(THT) = -1 \quad X(HTT) = -1 \quad X(TTT) = -3$$

## Number of pips in two dice.

“What is the likelihood of getting  $n$  pips?”



$$Pr[X = 10] = 3/36 = Pr[X^{-1}(10)]; Pr[X = 8] = 5/36 = Pr[X^{-1}(8)].$$

## Flip three coins

Experiment: flip three coins

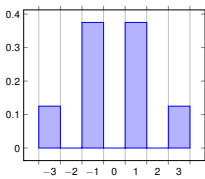
Sample Space:  $\{HHH, THH, HTH, TTH, HHT, THT, HTT, TTT\}$

Winnings: if win 1 on heads, lose 1 on tails.  $X$

Random Variable:  $\{3, 1, 1, -1, 1, -1, -1, -3\}$

Distribution:

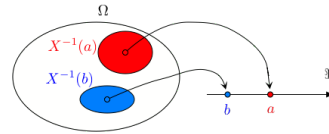
$$X = \begin{cases} -3, & \text{w. p. } 1/8 \\ -1, & \text{w. p. } 3/8 \\ 1, & \text{w. p. } 3/8 \\ 3, & \text{w. p. } 1/8 \end{cases}$$



## Distribution

The probability of  $X$  taking on a value  $a$ .

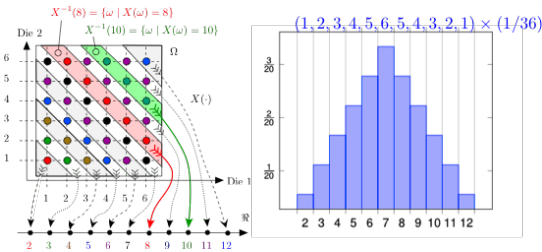
**Definition:** The **distribution** of a random variable  $X$ , is  $\{(a, Pr[X = a]) : a \in \mathcal{A}\}$ , where  $\mathcal{A}$  is the range of  $X$ .



$$Pr[X = a] := Pr[X^{-1}(a)] \text{ where } X^{-1}(a) := \{\omega \mid X(\omega) = a\}.$$

## Number of pips.

Experiment: roll two dice.



## Handing back assignments

Experiment: hand back assignments to 3 students at random.

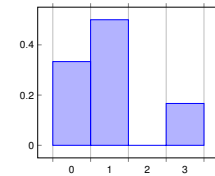
Sample Space:  $\Omega = \{123, 132, 213, 231, 312, 321\}$

How many students get back their own assignment?

Random Variable: values of  $X(\omega) : \{3, 1, 1, 0, 0, 1\}$

Distribution:

$$X = \begin{cases} 0, & \text{w.p. } 1/3 \\ 1, & \text{w.p. } 1/2 \\ 3, & \text{w.p. } 1/6 \end{cases}$$



## The binomial distribution.

Flip  $n$  coins with heads probability  $p$ .

Random variable: number of heads.

**Binomial Distribution:**  $Pr[X = i]$ , for each  $i$ .

How many sample points in event “ $X = i$ ”?

$i$  heads out of  $n$  coin flips  $\implies \binom{n}{i}$

What is the probability of  $\omega$  if  $\omega$  has  $i$  heads?

Probability of heads in any position is  $p$ .

Probability of tails in any position is  $(1 - p)$ .

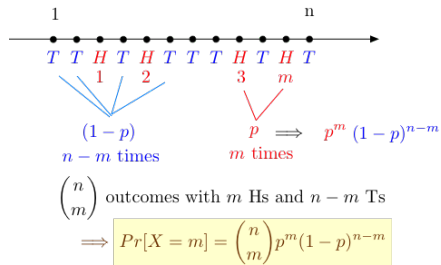
So, we get

$$Pr[\omega] = p^i(1-p)^{n-i}.$$

Probability of “ $X = i$ ” is sum of  $Pr[\omega]$ ,  $\omega \in “X = i”$ .

$$Pr[X = i] = \binom{n}{i} p^i(1-p)^{n-i}, i = 0, 1, \dots, n : B(n, p) \text{ distribution}$$

## The binomial distribution.



## Error channel.

A packet is corrupted with probability  $p$ .

Send  $n+2k$  packets.

Probability of at most  $k$  corruptions.

$$\sum_{i \leq k} \binom{n+2k}{i} p^i (1-p)^{n+2k-i}$$

Confidence in polling, experiments, etc.

## Combining Random Variables.

Let  $X$  and  $Y$  be two RV on the same probability space.

That is,  $X: \Omega \rightarrow \mathfrak{R}$  assigns the value  $X(\omega)$  to  $\omega$ . Also,  $Y: \Omega \rightarrow \mathfrak{R}$  assigns the value  $Y(\omega)$  to  $\omega$ .

Then  $X+Y$  is a random variable: It assigns the value

$$X(\omega) + Y(\omega)$$

to  $\omega$ .

Experiment: Roll two dice.  $X$  = outcome of first die,  $Y$  = outcome of second die. Thus,

$$X(a, b) = a \text{ and } Y(a, b) = b \text{ for } (a, b) \in \Omega = \{1, \dots, 6\}^2$$

Then  $Z = X + Y$  = sum of two dice is defined by

$$Z(a, b) = X(a, b) + Y(a, b) = a + b$$

## Combining Random Variables

Other random variables:

- ▶  $X^k: \Omega \rightarrow \mathfrak{R}$  is defined by  $X^k(\omega) = [X(\omega)]^k$ .  
In the dice example,  $X^3(a, b) = a^3$ .
- ▶  $(X-2)^2 + 4XY$  assigns the value  $(X(\omega)-2)^2 + 4X(\omega)Y(\omega)$  to  $\omega$ .
- ▶  $g(X, Y, Z)$  assigned the value  $g(X(\omega), Y(\omega), Z(\omega))$  to  $\omega$ .

## Expectation.

How did people do on the midterm?

Distribution.

Summary of distribution?

Average!



## Expectation - Intuition

Flip a loaded coin with  $Pr[H] = p$  a large number  $N$  of times.

We expect heads to come up a fraction  $p$  of the times and tails a fraction  $1-p$ .

Say that you get 5 for every  $H$  and 3 for every  $T$ .

If there are  $N(H)$  outcomes equal to  $H$  and  $N(T)$  outcomes equal to  $T$ , you collect

$$5 \times N(H) + 3 \times N(T)$$

because your average gain per experiment is then

$$\frac{5N(H) + 3N(T)}{N}$$

Since  $\frac{N(H)}{N} \approx p = Pr[X=5]$  and  $\frac{N(T)}{N} \approx 1-p = Pr[X=3]$ , we find that the average gain per outcome is approximately equal to

$$5Pr[X=5] + 3Pr[X=3]$$

We use this frequentist [interpretation](#) as a definition.

## Expectation - Definition

**Definition:** The **expected value** of a random variable  $X$  is

$$E[X] = \sum_a a \times Pr[X = a].$$

The expected value is also called the mean.

According to our intuition, we expect that if we repeat an experiment a large number  $N$  of times and if  $X_1, \dots, X_N$  are the successive values of the random variable, then

$$\frac{X_1 + \dots + X_N}{N} \approx E[X].$$

That is indeed the case, in the same way that the fraction of times that  $X = x$  approaches  $Pr[X = x]$ .

This (nontrivial) result is called the **Law of Large Numbers**.

The subjectivist interpretation of  $E[X]$  is less obvious.

## Expectation and Average.

There are  $n$  students in the class;

$X(m)$  = score of student  $m$ , for  $m = 1, 2, \dots, n$ .

"Average score" of the  $n$  students: add scores and divide by  $n$ :

$$\text{Average} = \frac{X(1) + X(2) + \dots + X(n)}{n}.$$

Experiment: choose a student uniformly at random.

Uniform sample space:  $\Omega = \{1, 2, \dots, n\}$ ,  $Pr[\omega] = 1/n$ , for all  $\omega$ .

Random Variable: midterm score:  $X(\omega)$ .

Expectation:

$$E(X) = \sum_{\omega} X(\omega) Pr[\omega] = \sum_{\omega} X(\omega) \frac{1}{n}.$$

Hence,

$$\text{Average} = E(X).$$

This holds for a **uniform** probability space.

## Expectation: A Useful Fact

**Theorem:**

$$E[X] = \sum_{\omega} X(\omega) \times Pr[\omega].$$

**Proof:**

$$\begin{aligned} E[X] &= \sum_a a \times Pr[X = a] \\ &= \sum_a a \times \sum_{\omega: X(\omega)=a} Pr[\omega] \\ &= \sum_a \sum_{\omega: X(\omega)=a} X(\omega) Pr[\omega] \\ &= \sum_{\omega} X(\omega) Pr[\omega] \end{aligned}$$

## Handing back assignments

We give back assignments randomly to three students.

What is the expected number of students that get their own assignment back?

"The expected number of **fixed points** in a random permutation."

Expected value of a random variable:

$$E[X] = \sum_a a \times Pr[X = a].$$

For 3 students (permutations of 3 elements):

$$Pr[X = 3] = 1/6, Pr[X = 1] = 1/2, Pr[X = 0] = 1/3.$$

$$E[X] = 3 \times \frac{1}{6} + 1 \times \frac{1}{2} + 0 \times \frac{1}{3} = 1.$$

## An Example

Flip a fair coin three times.

$\Omega = \{HHH, HHT, HTH, THH, HTT, THT, TTH, TTT\}$ .

$X$  = number of H's:  $\{3, 2, 2, 2, 1, 1, 1, 0\}$ .

Thus,

$$\sum_{\omega} X(\omega) Pr[\omega] = \{3 + 2 + 2 + 2 + 1 + 1 + 1 + 0\} \times \frac{1}{8}.$$

Also,

$$\sum_a a \times Pr[X = a] = 3 \times \frac{1}{8} + 2 \times \frac{3}{8} + 1 \times \frac{3}{8} + 0 \times \frac{1}{8}.$$

## Win or Lose.

Expected winnings for heads/tails games, with 3 flips?

$$E[X] = 3 \times \frac{1}{8} + 1 \times \frac{3}{8} - 1 \times \frac{3}{8} - 3 \times \frac{1}{8} = 0.$$

Can you ever win 0?

Apparently: expected value is not a common value, by any means.

## Expectation

Recall:  $X : \Omega \rightarrow \mathfrak{R}; Pr[X = a] = Pr[X^{-1}(a)]$ ;

**Definition:** The **expectation** of a random variable  $X$  is

$$E[X] = \sum_a a \times Pr[X = a].$$

**Indicator:**

Let  $A$  be an event. The random variable  $X$  defined by

$$X(\omega) = \begin{cases} 1, & \text{if } \omega \in A \\ 0, & \text{if } \omega \notin A \end{cases}$$

is called the **indicator** of the event  $A$ .

Note that  $Pr[X = 1] = Pr[A]$  and  $Pr[X = 0] = 1 - Pr[A]$ .

Hence,

$$E[X] = 1 \times Pr[X = 1] + 0 \times Pr[X = 0] = Pr[A].$$

The random variable  $X$  is sometimes written as

$$1_{\{\omega \in A\}} \text{ or } 1_A(\omega).$$

## Using Linearity - 2: Fixed point.

Hand out assignments at random to  $n$  students.

$X$  = number of students that get their own assignment back.

$X = X_1 + \dots + X_n$  where

$X_m = 1_{\{\text{student } m \text{ gets his/her own assignment back}\}}$ .

One has

$$\begin{aligned} E[X] &= E[X_1 + \dots + X_n] \\ &= E[X_1] + \dots + E[X_n], \text{ by linearity} \\ &= nE[X_1], \text{ because all the } X_m \text{ have the same distribution} \\ &= nPr[X_1 = 1], \text{ because } X_1 \text{ is an indicator} \\ &= n(1/n), \text{ because student 1 is equally likely} \\ &\quad \text{to get any one of the } n \text{ assignments} \\ &= 1. \end{aligned}$$

Note that linearity holds even though the  $X_m$  are not independent (whatever that means).

## Linearity of Expectation

**Theorem:**

$$E[X] = \sum_{\omega} X(\omega) \times Pr[\omega].$$

**Theorem:** Expectation is linear

$$E[a_1 X_1 + \dots + a_n X_n] = a_1 E[X_1] + \dots + a_n E[X_n].$$

**Proof:**

$$\begin{aligned} E[a_1 X_1 + \dots + a_n X_n] &= \sum_{\omega} (a_1 X_1 + \dots + a_n X_n)(\omega) Pr[\omega] \\ &= \sum_{\omega} (a_1 X_1(\omega) + \dots + a_n X_n(\omega)) Pr[\omega] \\ &= a_1 \sum_{\omega} X_1(\omega) Pr[\omega] + \dots + a_n \sum_{\omega} X_n(\omega) Pr[\omega] \\ &= a_1 E[X_1] + \dots + a_n E[X_n]. \end{aligned}$$

□

## Using Linearity - 3: Binomial Distribution.

Flip  $n$  coins with heads probability  $p$ .  $X$  = number of heads

**Binomial Distribution:**  $Pr[X = i]$ , for each  $i$ .

$$Pr[X = i] = \binom{n}{i} p^i (1-p)^{n-i}.$$

$$E[X] = \sum_i i \times Pr[X = i] = \sum_i i \times \binom{n}{i} p^i (1-p)^{n-i}.$$

Uh oh. ... Or... a better approach: Let

$$X_i = \begin{cases} 1 & \text{if } i\text{th flip is heads} \\ 0 & \text{otherwise} \end{cases}$$

$$E[X_i] = 1 \times Pr[\text{"heads"}] + 0 \times Pr[\text{"tails"}] = p.$$

Moreover  $X = X_1 + \dots + X_n$  and

$$E[X] = E[X_1] + E[X_2] + \dots + E[X_n] = n \times E[X_i] = np.$$

## Using Linearity - 1: Pips on dice

Roll a die  $n$  times.

$X_m$  = number of pips on roll  $m$ .

$X = X_1 + \dots + X_n$  = total number of pips in  $n$  rolls.

$$\begin{aligned} E[X] &= E[X_1 + \dots + X_n] \\ &= E[X_1] + \dots + E[X_n], \text{ by linearity} \\ &= nE[X_1], \text{ because the } X_m \text{ have the same distribution} \end{aligned}$$

Now,

$$E[X_1] = 1 \times \frac{1}{6} + \dots + 6 \times \frac{1}{6} = \frac{6 \times 7}{2} \times \frac{1}{6} = \frac{7}{2}.$$

Hence,

$$E[X] = \frac{7n}{2}.$$

## Summary

### Random Variables

- ▶ A random variable  $X$  is a function  $X : \Omega \rightarrow \mathfrak{R}$ .
- ▶  $Pr[X = a] := Pr[X^{-1}(a)] = Pr[\{\omega \mid X(\omega) = a\}]$ .
- ▶  $Pr[X \in A] := Pr[X^{-1}(A)]$ .
- ▶ The distribution of  $X$  is the list of possible values and their probability:  $\{(a, Pr[X = a]), a \in \mathcal{A}\}$ .
- ▶  $g(X, Y, Z)$  assigns the value ....
- ▶  $E[X] := \sum_a a Pr[X = a]$ .
- ▶ Expectation is Linear.
- ▶  $B(n, p)$ .

## Probability: Midterm 2 Review.

### ► Framework:

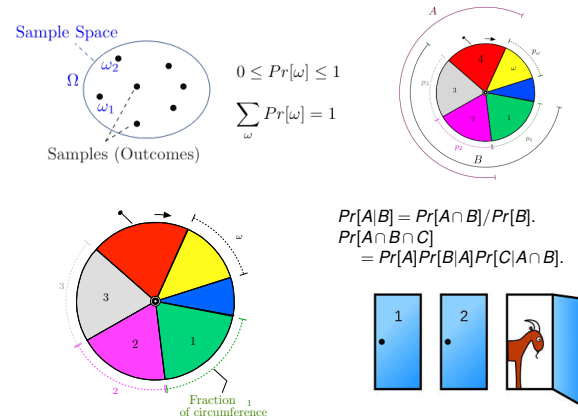
- Probability Space
- Conditional Probability & Bayes' Rule
- Independence
- Mutual Independence

### ► Collisions & Collecting

### ► Random Variables

See Note 25: 1, 2, 3, 4 (paragraphs 1, 2, 3; examples 1 through 8)

## Probability Space

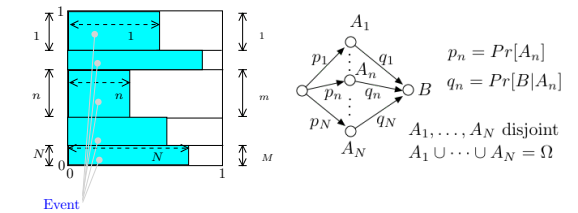


## Bayes' Rule

► Priors:  $Pr[A_n] = p_n, n = 1, \dots, M$

► Conditional Probabilities:  $Pr[B|A_n] = q_n, n = 1, \dots, N$

►  $\Rightarrow$  Posteriors:  $Pr[A_n|B] = \frac{p_n q_n}{p_1 q_1 + \dots + p_N q_N}$



## Bayes' Rule: Examples

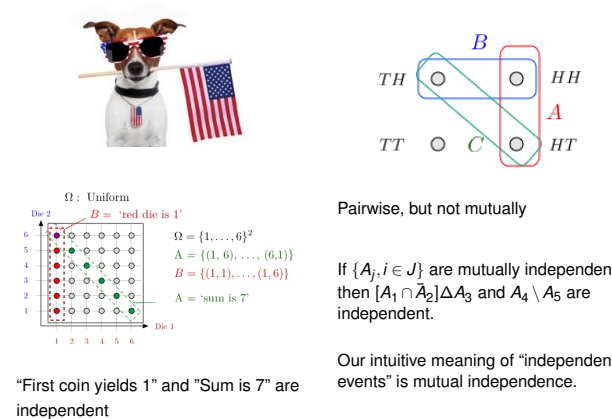
Let  $p'_n = Pr[A_n|B]$  be the posterior probabilities.

Thus,  $p'_n = p_n q_n / (p_1 q_1 + \dots + p_N q_N)$ .

Questions: Is it true that

- if  $q_n > q_k$ , then  $p'_n > p'_k$ ? Not necessarily.
- if  $p_n > p_k$ , then  $p'_n > p'_k$ ? Not necessarily.
- if  $p_n > p_k$  and  $q_n > q_k$ , then  $p'_n > p'_k$ ? Yes.
- if  $q_n = 1$ , then  $p'_n > 0$ ? Not necessarily.
- if  $p_n = 1/N$  for all  $n$ , then MLE = MAP? Yes.

## Independence



## Independence

Recall

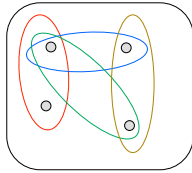
- $A$  and  $B$  are independent if  $Pr[A \cap B] = Pr[A]Pr[B]$ .
- $\{A_j, j \in J\}$  are mutually independent if  $Pr[\cap_{j \in K} A_j] = \prod_{j \in K} Pr[A_j], \forall$  finite  $K \subset J$ .

Thus,  $A, B, C, D$  are mutually independent if there are

- independent 2 by 2:  
 $Pr[A \cap B] = Pr[A]Pr[B], \dots, Pr[C \cap D] = Pr[C]Pr[D]$
- by 3:  $Pr[A \cap B \cap C] = Pr[A]Pr[B]Pr[C], \dots, Pr[B \cap C \cap D] = Pr[B]Pr[C]Pr[D]$
- by 4:  $Pr[A \cap B \cap C \cap D] = Pr[A]Pr[B]Pr[C]Pr[D]$ .

### Independence: Question 1

Consider the uniform probability space and the events  $A, B, C, D$ .



Which maximal collections of events among  $A, B, C, D$  are pairwise independent?

$\{A, B, C\}$ , and  $\{B, C, D\}$

Can you find three events among  $A, B, C, D$  that are mutually independent?

No: We would need an outcome with probability  $1/8$ .

### Independence: Question 2

Let  $\Omega = \{1, 2, \dots, p\}$  be a uniform probability space where  $p$  is prime.

Can you find two independent events  $A$  and  $B$  with  $Pr[A], Pr[B] \in (0, 1)$ ?

Let  $a = |A|, b = |B|, c = |A \cap B|$ .

Then,

$$Pr[A \cap B] = Pr[A]Pr[B], \text{ so that}$$

$$\frac{c}{p} = \frac{a}{p} \times \frac{b}{p}. \text{ Hence,}$$

$$ab = cp.$$

This is not possible since  $a, b < p$ .

### Collisions & Collecting

Collisions:

$$Pr[\text{no collision}] \approx e^{-m^2/2n}$$

Collecting:

$$Pr[\text{miss Wilson}] \approx e^{-m/n}$$

$$Pr[\text{miss at least one}] \leq ne^{-m/n}$$

### Math Tricks

Approximations:

$$\ln(1 - \varepsilon) \approx -\varepsilon$$

$$\exp\{-\varepsilon\} \approx 1 - \varepsilon$$

Sums:

$$(a+b)^n = \sum_{m=0}^n \binom{n}{m} a^m b^{n-m}$$

$$1 + 2 + \dots + n = \frac{n(n+1)}{2};$$

### Math Tricks, continued

Symmetry: E.g., if we pick balls from a bag, with no replacement,

$$Pr[\text{ball 5 is red}] = Pr[\text{ball 1 is red}]$$

Order of balls = permutation.

All permutations have same probability.

Union Bound:

$$Pr[A \cup B \cup C] \leq Pr[A] + Pr[B] + Pr[C]$$

Inclusion/Exclusion:

$$Pr[A \cup B] = Pr[A] + Pr[B] - Pr[A \cap B]$$

Total Probability:

$$Pr[B] = Pr[A_1]Pr[B|A_1] + \dots + Pr[A_n]Pr[B|A_n]$$

An  $L^2$ -bounded martingale converges almost surely. Just kidding!

### A mini-quiz

True or False:

- ▶  $Pr[A \cup B] = Pr[A] + Pr[B]$ . **False** True iff disjoint.
- ▶  $Pr[A \cap B] = Pr[A]Pr[B]$ . **False** True iff independent.
- ▶  $A \cap B = \emptyset \Rightarrow A, B$  independent. **False**
- ▶ For all  $A, B$ , one has  $Pr[A|B] \geq Pr[A]$ . **False**
- ▶  $Pr[A \cap B \cap C] = Pr[A]Pr[B|A]Pr[C|B]$ . **False**

## A mini-quizz; part 2

- ▶  $\Omega = \{1, 2, 3, 4\}$ , uniform. Find events  $A, B, C$  that are pairwise independent, not mutually.  
 $A = \{1, 2\}, B = \{1, 3\}, C = \{1, 4\}$ .
- ▶  $A, B, C$  pairwise independent. Is it true that  $(A \cap B)$  and  $C$  are independent?  
No. In example above,  $Pr[A \cap B \cap C] \neq Pr[A \cap B]Pr[C]$ .
- ▶ Assume  $Pr[C|A] > Pr[C|B]$ .  
Is it true that  $Pr[A|C] > Pr[B|C]$ ?  
No.
- ▶ Deal two cards from a 52-card deck. What is the probability that the value of the first card is strictly larger than that of the second?  
 $Pr[\text{same}] = \frac{3}{51}$ .  $Pr[\text{different}] = \frac{48}{51}$ .  
 $Pr[\text{first} > \text{second}] = \frac{24}{51}$ .

## Summary

Good clean fun ....

And good time was had by all ....

Enjoy spring break and the midterm.