

# EECS 126: Probability & Random Processes

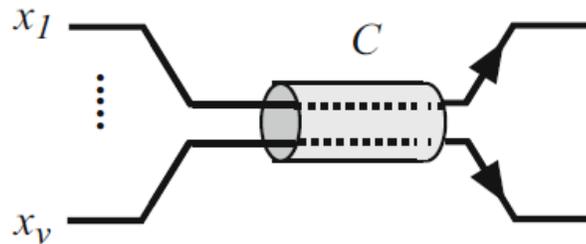
## Fall 2021

Multiplexing

Shyam Parekh

## Multiplexing

- Canonical Example: Sharing of a link by multiple connections.

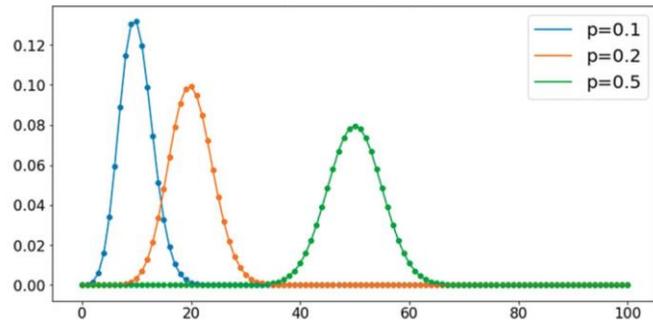


- Each connection will get the rate of  $C/v$ , where the RV  $v$  represents # of active connections.
- In general, multiplexing addresses sharing of a common resource.
- We'll consider different statistical aspects of multiplexing.

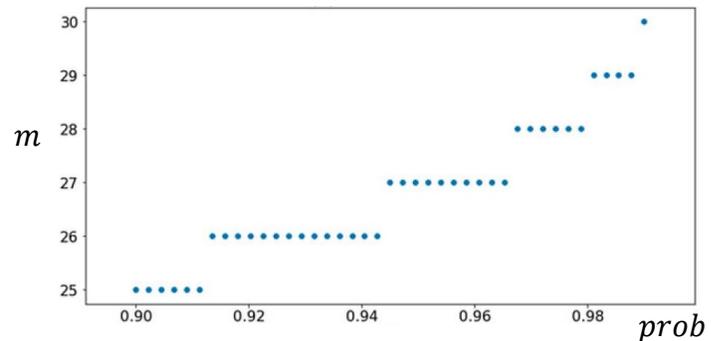
## Binomial Distribution

- Consider the link sharing example.
- $v \equiv_D B(N, p)$  can be used to model the number of active users.

Probability Mass Function  $B(100, p)$



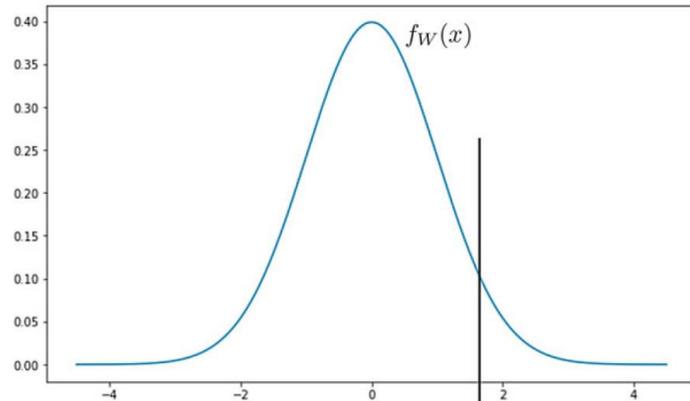
Percent Point Function (CDF Inverse) of  $B(100, 0.2)$



- Find smallest  $m$  s.t.  $P(v > m) \leq 0.05$ , or equivalently  $P(v \leq m) > 0.95$ 
  - This would imply each active user will get at least a rate of  $C/m$  with probability 95% or higher.
  - Observe  $P(v \leq 27) = 0.966 > 0.95$  and  $P(v \leq 26) = 0.944 < 0.95$ .

## Gaussian Random Variable

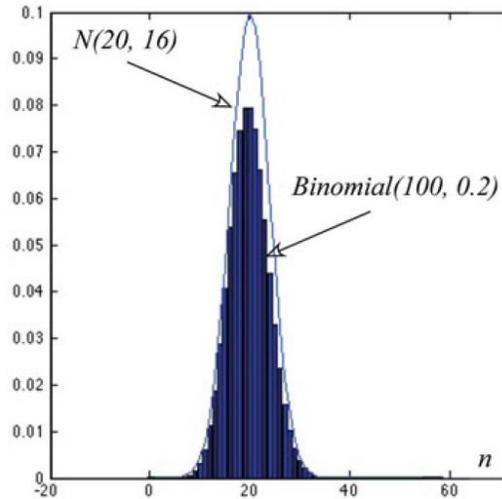
- $X \stackrel{D}{=} \mathcal{N}(\mu, \sigma^2)$ 
  - PDF  $f_X(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2\right)$ ,  $-\infty < x < \infty$ .
  - Note  $X = \mu + \sigma W$ , where  $W \stackrel{D}{=} \mathcal{N}(0, 1)$  is a standard normal/Gaussian RV.



- Useful facts:  $P(W > 1.65) \approx 0.05$ ,  $P(W > 1.96) \approx 0.025$ ,  $P(W > 2.32) \approx 0.01$ .
- Above facts hold for general Gaussian RV larger than the mean by a scalar multiple of the standard deviation (e.g.,  $P(X > \mu + 1.65\sigma) = 0.05$ , where  $X \stackrel{D}{=} \mathcal{N}(\mu, \sigma^2)$ ).

## Central Limit Theorem (CLT)

- Convergence in Distribution: Let  $\{X(n), n \geq 1\}$  and  $X$  be random variables. We say  $X(n)$  converges in distribution to  $X$ , and write  $X(n) \Rightarrow X$ , if  $P(X(n) \leq x) \rightarrow P(X \leq x) \forall x$  s. t.  $P(X = x) = 0$ .
  - $X(n) \rightarrow X$  a. s. implies  $X(n) \rightarrow X$  in probability implies  $X(n) \Rightarrow X$ . (See proofs [here](#).)
- CLT Theorem: Let  $\{X(n), n \geq 1\}$  be IID RVs with mean  $E(X(n)) = \mu$ , and  $var(X(n)) = \sigma^2$ . Then, as  $n \rightarrow \infty$ ,  $\frac{X(1)+X(2)+\dots+X(n)-n\mu}{\sigma\sqrt{n}} \Rightarrow \mathcal{N}(0, 1)$ .
- Binomial and Gaussian:
  - BY CLT,  $B(N, p) \approx \mathcal{N}(Np, Np(1-p))$ .



- $P(B(N, p) > Np + 1.65\sqrt{Np(1-p)}) \approx 0.05$ . Hence  $m = Np + 1.65\sqrt{Np(1-p)} \approx 27$  for  $B(100, 0.2)$ .

## Confidence Intervals

- Let  $Y(N) \equiv_D B(N, p)/N$ , and define  $A_1 = \{Y(N) \geq p + 1.65\sqrt{p(1-p)/N}\}$  and  $A_2 = \{Y(N) \leq p - 1.65\sqrt{p(1-p)/N}\}$ .
- Due to CLT,  $P(A_1 \cup A_2) \approx 0.1$ , or  $P(A_1^c \cap A_2^c) \approx 0.9$ .
- Since  $p(1-p) < 1/4$ , we have  $P\left(Y(N) - 0.83\frac{1}{\sqrt{N}} \leq p \leq Y(N) + 0.83\frac{1}{\sqrt{N}}\right) \geq 0.9$ .
  - 90% Confidence Interval.
  - Replacing 1.65 by 2 (i.e., 0.83 by 1), gives 95% Confidence Interval.
- Increasing  $N$ , shrinks the interval.
  - Determine  $N$  based on the desired estimation margin.
- Consider any IID RVs  $\{X(n), n \geq 1\}$  with mean  $\mu$  without knowledge of a bound on the variance.
  - Let the sample mean  $\mu_n = \frac{X(1)+X(2)+\dots+X(n)}{n}$ .
  - Let  $\sigma_n$  be the sample standard deviation, where  $\sigma_n^2 = \frac{\sum_{m=1}^n (X(m) - \mu_n)^2}{n-1}$ .
  - 90% Confidence Interval for  $\mu$ :  $\left[\mu_n - 1.65 \frac{\sigma_n}{\sqrt{n}}, \mu_n + 1.65 \frac{\sigma_n}{\sqrt{n}}\right]$ .
  - 95% Confidence Interval for  $\mu$ :  $\left[\mu_n - 2 \frac{\sigma_n}{\sqrt{n}}, \mu_n + 2 \frac{\sigma_n}{\sqrt{n}}\right]$ .

## Characteristic Functions

- Definition: The characteristic function of a random variable  $X$  is defined as  $\phi_X(u) = E(e^{iuX})$ ,  $u \in \mathcal{R}$ , where  $i = \sqrt{-1}$ .
  - It's similar to the Moment Generating Function (MGF)  $M_X(t) := E(e^{tX})$ ,  $t \in \mathcal{R}$ .
  - Let  $X \equiv_D \mathcal{N}(0, 1)$ . Then,  $\phi_X(u) = e^{-\frac{u^2}{2}}$ .
  - A characteristic function determines the associated PDF/CDF uniquely.
- Moments of  $\mathcal{N}(0, 1)$ .

## Proof of Central Limit Theorem (Sketch)

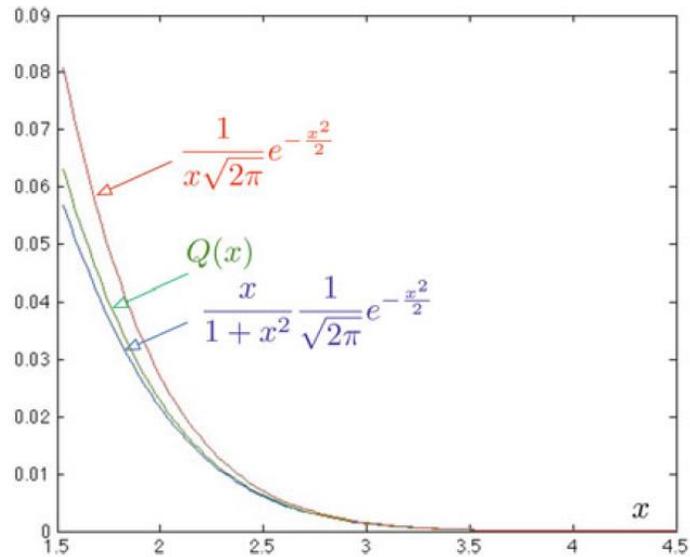
- Theorem: Let  $\{X(n), n \geq 1\}$  be IID RVs with mean  $E(X(n)) = \mu$ , and  $\text{var}(X(n)) = \sigma^2$ . Then, as  $n \rightarrow \infty$ ,  $\frac{X(1)+X(2)+\dots+X(n)-n\mu}{\sigma\sqrt{n}} \Rightarrow \mathcal{N}(0, 1)$ .

## Two Applications of Characteristic Functions

- Poisson as a limit of Binomial:  $B(n, \lambda/n) \Rightarrow P(\lambda)$ .
- Exponential as a limit of Geometric:  $G(\lambda/n)/n \Rightarrow \text{Exp}(\lambda)$ .

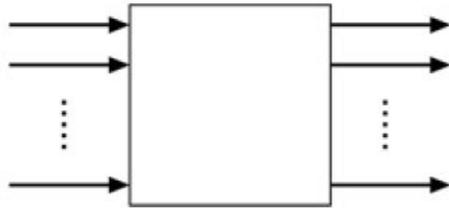
## Miscellaneous Results

- Let  $X, Y$  be IID  $\mathcal{N}(0, 1)$  RVs. Then,  $Z := X^2 + Y^2 \equiv_D \text{Exp}(1/2)$ .
- Let  $Q(x) := P(X > x)$ ,  $X \equiv_D \mathcal{N}(0, 1)$ . Then,  $\frac{x}{1+x^2} f_X(x) \leq Q(x) \leq \frac{1}{x} f_X(x)$ ,  $\forall x > 0$ , where  $f_X(x)$  is the PDF of  $X$ .



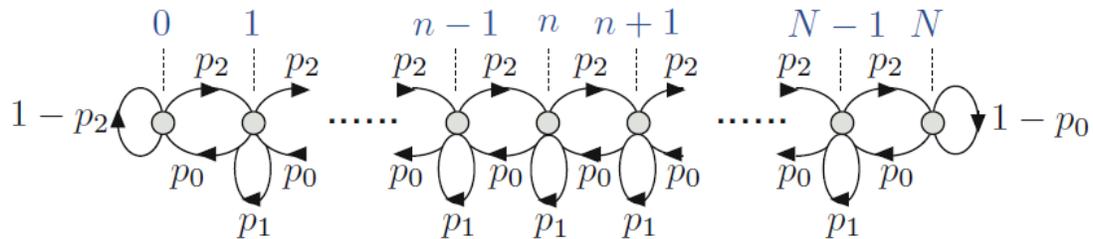
## Buffers at a switch

- Multiplexing of traffic at an output port of a switch can cause buffering.



- DTMC model for buffering at an output port.
- At each instant  $n+$  (right after time  $n$ )\* a packet arrives with probability  $\lambda \in [0, 1]$  independently of the past.
- Time to transmit a packet is geometrically distributed with parameter  $\mu \in (0, 1]$ , all transmission times are independent.
  - A packet in service completes transmission at  $n-$  (just before time  $n$ )\* with probability  $\mu$ .
- Let  $\{X_n, n \geq 0\}$  be the number of packets in the output buffer at time  $n$ .

$$p_0 = \mu(1 - \lambda), p_2 = \lambda(1 - \mu), p_1 = 1 - p_0 - p_2$$



\*This clarification of arrival/departure at  $n+/n-$  is not in the textbook.

## Buffers at a switch (Cont'd)

- DTMC Analysis:
  - Let  $N$  be the buffer capacity in terms of number of packets.
  - Balance equations yield the invariant distribution  $\pi$  with  $\pi(i) = \pi(0)\rho^i, i = 0, 1, \dots, N$ , where  $\rho = p_2/p_0$ , and  $\pi(0) = \frac{1-\rho}{1-\rho^{N+1}}$ .
  - Average # of packets in the buffer under the invariant distribution  $\approx \frac{\rho}{1-\rho}$  (assuming  $\lambda < \mu$  and  $N \gg 1$ ).
  - Average delay in the buffer (from arrival until service completion)  $\approx \frac{1-\mu}{\mu-\lambda}$ .
- Little's Law: Let  $L$  = average # in the system,  $\lambda$  = average arrival rate and  $W$  = average time in the system, then  $L = \lambda W$ .





