

## 1 Overview and Motivation

In this note, we will examine two important, connected, applications of the singular value decomposition (SVD), discussed in [Note 14](#).

The first application is using *low rank approximations* for dimensionality reduction of data.

### Key Idea 1 (Low-Rank Approximation)

*Low-rank approximation* of a matrix  $A \in \mathbb{R}^{m \times n}$  with rank  $r$  is the process of finding another matrix  $A_\ell \in \mathbb{R}^{m \times n}$  with rank  $\ell \ll r$  such that  $A - A_\ell$  is "small" in some sense.

Our second application is a tool for data analysis. When we collect data, there are potentially lots of factors that influence the values we see. Our goal is to figure out the most important such factors. This allows us to compress our data and remove the dimensions that are not important. We do this via a technique called *principal component analysis* (PCA).

### Key Idea 2 (Principal Component Analysis)

*Principal component analysis* is a way of capturing the most important dimensions of data.

## 2 Low-Rank Approximation

Given a matrix  $A \in \mathbb{R}^{m \times n}$  of rank  $r \leq \min\{m, n\}$ , we saw in [Note 14](#) that we can write  $A$  using the outer-product form of the SVD:

$$A = \sum_{i=1}^r \sigma_i \vec{u}_i \vec{v}_i^\top. \quad (1)$$

If our matrix is high-rank, i.e.,  $r \approx \min\{m, n\}$ , then almost all the  $\sigma_i$  will be nonzero and non-negligible. However, if the data has some linear, low-rank essential structure, as is usually the case with real data such as images, most of our singular values will be very small (but usually nonzero due to noise or disturbances). If, say, the data has intrinsic linear rank  $\ell$ , then the first  $\ell$  singular values are large, and the remaining  $r - \ell$  are small:

$$A = \sum_{i=1}^r \sigma_i \vec{u}_i \vec{v}_i^\top \quad (2)$$

$$= \sum_{i=1}^{\ell} \sigma_i \vec{u}_i \vec{v}_i^\top + \sum_{i=\ell+1}^r \underbrace{\sigma_i}_{\approx 0} \vec{u}_i \vec{v}_i^\top \quad (3)$$

$$\approx \sum_{i=1}^{\ell} \sigma_i \vec{u}_i \vec{v}_i^\top. \quad (4)$$

This motivates approximating the data as

$$A_\ell := \sum_{i=1}^{\ell} \sigma_i \vec{u}_i \vec{v}_i^\top \quad (5)$$

and using this compressed data for further analysis.

For notation's sake, if we define

$$U_\ell := [\vec{u}_1 \quad \cdots \quad \vec{u}_\ell] \quad V_\ell := [\vec{v}_1 \quad \cdots \quad \vec{v}_\ell] \quad \Sigma_\ell := \begin{bmatrix} \sigma_1 & & \\ & \ddots & \\ & & \sigma_\ell \end{bmatrix} \quad (6)$$

then  $A_\ell = U_\ell \Sigma_\ell V_\ell^\top$ . Under this notation  $A_r = U_r \Sigma_r V_r^\top$  is the compact SVD of  $A$ , so  $A_r = A$ .

The approximation of  $A$  by  $A_\ell$  is theoretically justified by the *Eckart-Young-Mirsky theorem* (sometimes just *Eckart-Young theorem*), which says that this is the best rank- $\ell$  approximation in terms of the Frobenius norm (recall [Note 8](#) and [Note 11](#)).<sup>1</sup>

### Theorem 3 (Eckart-Young-Mirsky Theorem)

Let  $A \in \mathbb{R}^{m \times n}$  have rank  $r \leq \min\{m, n\}$ . For  $\ell \leq r$  and  $A_\ell$  as defined above, we have that

$$A_\ell \in \operatorname{argmin}_{B \in \mathbb{R}^{m \times n}} \|A - B\|_F^2 \quad (7)$$

$$\text{s.t. } \operatorname{rank}(B) = \ell. \quad (8)$$

The proof of Theorem 3 is on the longer side and may distract from the overall flow of this note, so it is left to [Appendix A](#). For this iteration of the course, the proof is not in scope.

NOTE: The theorem is also true if we use the *relaxed* constraint  $\operatorname{rank}(B) \leq \ell$ .

## 3 Principal Component Analysis

Suppose we have collected some noisy data points, which are represented as vectors  $\vec{x}_1, \dots, \vec{x}_n \in \mathbb{R}^d$ . Let

$$A := [\vec{x}_1 \quad \cdots \quad \vec{x}_n], \quad (9)$$

be the so-called *data matrix*, i.e., arranging the data points as the *columns* of  $A$ . Suppose  $\operatorname{rank}(A) = r \leq \min\{n, d\}$ . Further suppose the so-called "ground truth" data (i.e., data without the noise) actually would span a  $\ell$ -dimensional subspace  $S_{\text{gt}}$  of  $\mathbb{R}^d$ , with  $\ell \leq r$ .<sup>2</sup> The remaining  $r - \ell$  dimensions, in this case, would be the product of noise. The goal of principal component analysis (PCA) is to find, from the noisy data, the relevant  $\ell$ -dimensional subspace  $S_{\text{gt}}$  which the dataset spans.

One general approach to solving problems of this type is:

1. Make an *objective function* which quantifies the property you want to optimize.
2. Optimize the objective function, either by hand or by computer (the latter is more popular nowadays).

<sup>1</sup>This is also true for some other norms, such as the *spectral norm*, which we are not covering in this class.

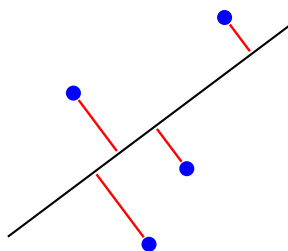
<sup>2</sup>In this context, we are assuming that the number  $\ell$  is known to us; later we will explain what to do if  $\ell$  is unknown, as is very often the case in practice.

This seems very abstract. Applying it to our problem, we see that we want to estimate the "best"  $\ell$ -dimensional subspace that our ground truth data would span. We would then ask what the meaning of "best" is; a reasonable and quantifiable notion of "best subspace" is "the one closest to all the points". From here we can develop an objective function that we can minimize. The objective function would take in a subspace  $S$ , and evaluate how close it is to all the points; in particular, it would be the following:

$$\sum_{i=1}^n \|\vec{x}_i - \text{proj}_S(\vec{x}_i)\|^2. \quad (10)$$

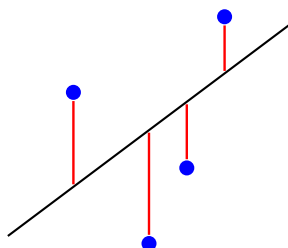
For some intuition, if  $n = 4$  and  $d = 2$  (i.e., our data set is 4 points in  $\mathbb{R}^2$ ), this objective function would be the sum of the squared lengths of the red lines in this picture, where the subspace  $S$  is a black line and the data points  $\vec{x}_1, \dots, \vec{x}_4$  are blue dots:

**Figure 1:** Objective function visualization for PCA.



Contrast this to the least squares objective function, which computes the sum of squares of the *vertical* residuals (instead of the orthogonal residuals):

**Figure 2:** Objective function visualization for least squares.



Now we explore how to compute this objective function. From [Note 11](#), we know that, if  $W \in \mathbb{R}^{d \times \ell}$  has orthonormal columns which span  $S$ , i.e.,  $W^\top W = I_\ell$  and  $\text{Col}(W) = S$ , then

$$\text{proj}_S(\vec{x}_i) = WW^\top \vec{x}_i. \quad (11)$$

This means that our objective function takes the form

$$\sum_{i=1}^n \|\vec{x}_i - WW^\top \vec{x}_i\|^2. \quad (12)$$

And we would try to minimize this over subspaces  $S$ , i.e., over matrices  $W \in \mathbb{R}^{d \times \ell}$  with orthonormal columns such that  $W^\top W = I_\ell$ . This leads to the following optimization problem, which can be efficiently solved via the SVD.

**Theorem 4** (Principal Component Analysis)

Let  $\vec{x}_1, \dots, \vec{x}_n \in \mathbb{R}^d$  be data points. If  $A = \begin{bmatrix} \vec{x}_1 & \dots & \vec{x}_n \end{bmatrix}$  is the data matrix, with SVD  $A = U\Sigma V^\top$ , then

$$U_\ell \in \operatorname{argmin}_{W \in \mathbb{R}^{d \times \ell}} \sum_{i=1}^n \left\| \vec{x}_i - WW^\top \vec{x}_i \right\|^2 \quad (13)$$

$$\text{s.t. } W^\top W = I_\ell. \quad (14)$$

where  $U_\ell = \begin{bmatrix} \vec{u}_1 & \dots & \vec{u}_\ell \end{bmatrix}$  is the first  $\ell$  columns of  $U$ .

The proof of Theorem 4 is on the longer side and may distract from the overall flow of this note, so it is left to Appendix B. We fully expect you to read the proof and understand it. It is completely in-scope for the course.

We may also phrase this result geometrically, in terms of subspaces instead of matrices.

**Corollary 5** (Geometric Principal Component Analysis)

Let  $\vec{x}_1, \dots, \vec{x}_n \in \mathbb{R}^d$  be data points. If  $A = \begin{bmatrix} \vec{x}_1 & \dots & \vec{x}_n \end{bmatrix}$  is the data matrix, with SVD  $A = U\Sigma V^\top$ , then

$$S_{\text{PCA}} \in \operatorname{argmin}_{S \subseteq \mathbb{R}^d} \sum_{i=1}^n \left\| \vec{x}_i - \operatorname{proj}_S(\vec{x}_i) \right\|^2 \quad (15)$$

$$\text{s.t. } \dim(S) \leq \ell \quad (16)$$

where  $S_{\text{PCA}} = \operatorname{Span}(\vec{u}_1, \dots, \vec{u}_\ell)$  is the span of the first  $\ell$  columns of  $U$ , and the argmin is taken over subspaces of  $\mathbb{R}^d$ .

In fact, these  $\vec{u}_i$  vectors are called the *principal components* of the data.

**Definition 6** (Principal Components)

Let  $A = \begin{bmatrix} \vec{x}_1 & \dots & \vec{x}_n \end{bmatrix} = U\Sigma V^\top$  be a data matrix. The *principal components* of  $A$  are the vectors  $\vec{u}_1, \dots, \vec{u}_d$  in that order. (For instance,  $\vec{u}_1$  is the first principal component,  $\vec{u}_2$  is the second principal component, etc.)

Based on our knowledge from Note 14 of how  $\vec{u}_i$ 's are the eigenvectors of  $AA^\top$ , we have the following equivalent, alternate characterization, which tells us how to calculate the principal components without calculating the SVD.

**Theorem 7** (Principal Components as Eigenvectors)

Let  $A = \begin{bmatrix} \vec{x}_1 & \dots & \vec{x}_n \end{bmatrix} \in \mathbb{R}^{d \times n}$  be a data matrix. The principal components of  $A$  are the eigenvectors of  $AA^\top$ , ordered in non-increasing order by the value of the corresponding eigenvalue, with ties broken arbitrarily.

Putting this into an algorithm gets us the following methods to find principal components, depending if the data points are columns or rows.

Here the  $\operatorname{SORT}(A, D)$  function sorts the columns of  $A$  in non-increasing order by the values on the diagonal of  $D$ , breaking ties arbitrarily.

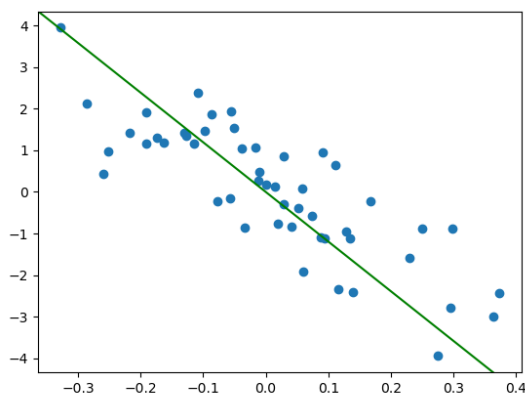
**Algorithm 8** Principal Component Analysis

- 
- 1: **function** FINDPRINCIPALCOMPONENTS( $A, \ell$ )
  - 2:    $(U, \Lambda) := \text{DIAGONALIZE}(AA^\top)$
  - 3:   **return**  $U_\ell :=$  the first  $\ell$  columns of SORT( $U, \Lambda$ )
  - 4: **end function**
- 

*NOTE:* In practice, one sometimes subtracts the column mean  $\frac{1}{n} \cdot A \vec{1}_n$  from each column of  $A$ , as a pre-processing step. That is, we replace  $A$  with  $A_{\text{new}} := A(I_n - \frac{1}{n} \vec{1}_n \vec{1}_n^\top)$ , then diagonalize  $A_{\text{new}} A_{\text{new}}^\top$ , and take its eigenvectors to be the principal components. The precise reason why is left to a statistics course.

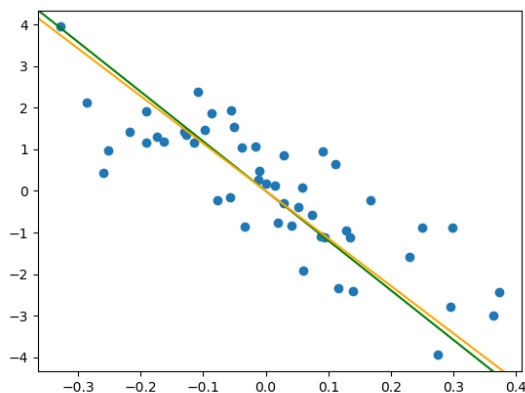
To give an idea of how effective this procedure is, suppose we have the following data (blue), with the "ground truth" subspace  $S_{\text{gt}}$  (green) also shown.

**Figure 3:** A sample dataset with  $n = 50$ ,  $d = 2$ , and  $\dim(S_{\text{gt}}) = 1$  (green).



If we also plot the first principal component subspace (shown in orange), we get the following plot.

**Figure 4:** A sample dataset with  $n = 50$ ,  $d = 2$ ,  $\dim(S_{\text{gt}}) = 1$  (green), and  $\dim(S_{\text{PCA}}) = 1$  (orange).



The idea is: *the principal components learned from data, are very similar to the true low-rank subspace the ground truth data lie on!* This idea generalizes to many dimensions as well.

See Appendix C for the code that generates these plots.

### Warning 9

All algorithms in this section work for *column data* – if we are working with *row data* of the form

$A = \begin{bmatrix} \vec{x}_1^\top \\ \vdots \\ \vec{x}_n^\top \end{bmatrix} \in \mathbb{R}^{n \times d}$ , which is very common in modern machine learning applications, we should take the transpose of the data matrix and then do PCA using the algorithms we already talked about.

## 4 Denoising and Dimensionality Reduction

Suppose we have a dataset  $A \in \mathbb{R}^{d \times n}$  (where the data points are columns of  $A$ ) which we perform PCA on, and get  $\ell$  principal components  $\vec{u}_1, \dots, \vec{u}_\ell$ . Suppose we are given a new noisy vector  $\vec{x} \in \mathbb{R}^d$ , and we know that the "ground truth" de-noised vector approximately lies in our low-rank  $\ell$ -dimensional subspace. To recover an estimate for the original ground truth vector, we project onto this subspace, i.e., our principal component subspace. From what we know about projections (see Note 11 again) the formula to project onto  $\text{Span}(\vec{u}_1, \dots, \vec{u}_\ell) = \text{Col}(U_\ell)$  is

$$\text{proj}_{\text{Span}(\vec{u}_1, \dots, \vec{u}_\ell)}(\vec{x}) = \text{proj}_{\text{Col}(U_\ell)}(\vec{x}) = U_\ell U_\ell^\top \vec{x}. \quad (17)$$

Since this process removes the residual noise and preserves the essential low-rank structure of  $\vec{x}$ , we call it *denoising*. This yields another formal algorithm:

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### Algorithm 10 Denoising using PCA

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- 1: **function** PCADENOISING( $A, \vec{x}, \ell$ )
  - 2:      $U_\ell := \text{FINDPRINCIPALCOMPONENTS}(A, \ell)$
  - 3:     **return**  $\hat{x} := U_\ell U_\ell^\top \vec{x}$
  - 4: **end function**
- 

We may also discuss PCA from the viewpoint of *data compression* and *dimensionality reduction*. Namely, if  $\ell \leq d$ , then we can approximately represent points  $\vec{x}$  in  $\mathbb{R}^d$  in our dataset, by vectors  $\vec{w}$  in  $\mathbb{R}^\ell$ ; these vectors  $\vec{w}$  in  $\mathbb{R}^\ell$  precisely store the coefficients in the linear combination of  $\vec{u}_1, \dots, \vec{u}_\ell$  required to approximately generate  $\vec{x}$ . The closest point in the  $\ell$ -dimensional principal component subspace to  $\vec{x}$  is its projection, and again from Note 11 we know it has the formula

$$\text{proj}_{\text{Span}(\vec{u}_1, \dots, \vec{u}_\ell)}(\vec{x}) = U_\ell U_\ell^\top \vec{x} = \sum_{i=1}^{\ell} \langle \vec{x}, \vec{u}_i \rangle \vec{u}_i. \quad (18)$$

Then the coefficients  $w_i$  which make up our compressed vector  $\vec{w}$  are precisely the inner products  $\langle \vec{x}, \vec{u}_i \rangle$ , so we have

$$\vec{w} = \begin{bmatrix} w_1 \\ \vdots \\ w_\ell \end{bmatrix} = \begin{bmatrix} \langle \vec{x}, \vec{u}_1 \rangle \\ \vdots \\ \langle \vec{x}, \vec{u}_\ell \rangle \end{bmatrix} = \begin{bmatrix} \vec{u}_1^\top \vec{x} \\ \vdots \\ \vec{u}_\ell^\top \vec{x} \end{bmatrix} = \begin{bmatrix} \vec{u}_1^\top \\ \vdots \\ \vec{u}_\ell^\top \end{bmatrix} \vec{x} = U_\ell^\top \vec{x}. \quad (19)$$

To summarize, we can represent a vector  $\vec{x} \in \mathbb{R}^d$  in our  $\ell$ -dimensional subspace by  $\ell$  coordinates – that is a vector  $U_\ell^\top \vec{x}$  in  $\mathbb{R}^\ell$  whose entries are the inner products of  $\vec{x}$  with the first  $\ell$  principal components. Even if  $\vec{x}$

is not in our  $\ell$ -dimensional subspace, we can represent  $\vec{x}$  by a vector in  $\mathbb{R}^\ell$  which approximately generates  $\vec{x}$  (as opposed to exactly generating  $\vec{x}$ ). To recover the original vector (or in the latter case, an estimate for it), we can just multiply by  $U_\ell$  again, which reduces to Algorithm 10.

If  $\ell \ll d$  then this is a big success for us, since we have compressed our data a lot, and distilled it to its most important directions of variation. This is useful for data analysis, both from a computational efficiency standpoint and also an algorithm power standpoint; the reason why is out of scope of the course.

This yields another formal algorithm.

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**Algorithm 11** Dimensionality Reduction using PCA
 

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1: function PCADIMENSIONALITYREDUCTION( $A, \vec{x}, \ell$ )
2:    $U_\ell :=$  FINDPRINCIPALCOMPONENTS( $A, \ell$ )
3:   return  $\vec{w} := U_\ell^\top \vec{x}$ 
4: end function
  
```

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**Warning 12**

All algorithms in this section work for *column data* – if we are working with *row data* of the form

$A = \begin{bmatrix} \vec{x}_1^\top \\ \vdots \\ \vec{x}_n^\top \end{bmatrix} \in \mathbb{R}^{n \times d}$ , which is very common in modern machine learning applications, we should take

the transpose of the data matrix and then do PCA using the algorithms we already talked about.

## 5 Picking the Best Number of Principal Components

Throughout this note, we have generated a fixed number of principal components  $\ell$ . This assumes that we know our ground truth data has an  $\ell$ -dimensional low-rank structure, i.e., it lies on some  $\ell$ -dimensional subspace  $S_{\text{gt}}$ . However, in real-world applications, we would not know the dimensionality of the true  $S_{\text{gt}}$ . There are a few things we can do about this:

1. If we have hardware constraints, like in the lab, and can only take the first few principal components, then we should just use those.
2. If we are not constrained by hardware, then one thing we can do is the following:
  - (a) Separate our data into a training set  $\vec{x}_1, \dots, \vec{x}_{n_{\text{train}}}$  and a validation set  $\vec{x}_{1;\text{val}}, \dots, \vec{x}_{n_{\text{val}};\text{val}}$ .
  - (b) Obtain principal components using *only* the training data; i.e., find the eigenvectors of  $A_{\text{train}} A_{\text{train}}^\top$  where  $A_{\text{train}} := \begin{bmatrix} \vec{x}_1 & \cdots & \vec{x}_{n_{\text{train}}} \end{bmatrix}$ .
  - (c) Measure distance to the generated subspace using *only* the validation data, i.e., compute

$$\sum_{i=1}^{n_{\text{val}}} \left\| \vec{x}_{i;\text{val}} - U_k U_k^\top \vec{x}_{i;\text{val}} \right\|^2. \quad (20)$$

- (d) Stop taking additional components when this distance plateaus, i.e., stays the same after adding one or two more components. Because the principal component subspace grows larger with every iteration, our performance never becomes strictly worse with subsequent iterations, even

if we are using the validation set. So, we cannot stop when our performance gets strictly worse; this is the next best stopping rule.

This stops us from adding extraneous principal components which actually capture the effect of noise on the training set. Formally, this can be turned into an algorithm.

In the following algorithm, we include a "stopping parameter"  $\epsilon$ . This corresponds to a minimum improvement in the objective function that the new principal component needs to give, or else we conclude that we are done adding principal components and stop there.

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**Algorithm 13** A validation algorithm for PCA.

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```

function PCAVALIDATION( $A_{\text{train}}, A_{\text{val}}, \epsilon$ )
   $U := \text{FINDPRINCIPALCOMPONENTS}(A_{\text{train}}, d)$ 
   $s_0 = +\infty$  ▷ Objective function value
  for  $i \in \{1, \dots, d\}$  do
     $U_i :=$  first  $i$  columns of  $U$ 
    Compute  $s_i = \sum_{k=1}^{n_{\text{val}}} \|\vec{x}_{k,\text{val}} - U_i U_i^\top \vec{x}_{k,\text{val}}\|^2$ 
    if  $s_i - s_{i-1} \leq \epsilon$  then ▷ Stopping threshold
      return  $i$  ▷ Return the right number of principal components to use.
    end if
  end for
  return  $d$ 
end function

```

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#### Warning 14

All algorithms in this section work for *column data* – if we are working with *row data* of the form

$A = \begin{bmatrix} \vec{x}_1^\top \\ \vdots \\ \vec{x}_n^\top \end{bmatrix} \in \mathbb{R}^{n \times d}$ , which is very common in modern machine learning applications, we should take

the transpose of the data matrix and then do PCA using the algorithms we already talked about.

## 6 Examples

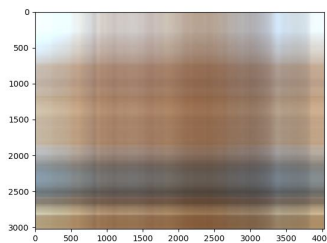
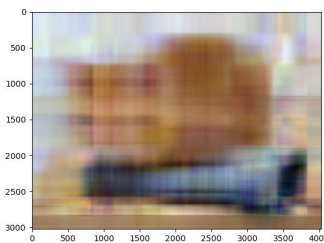
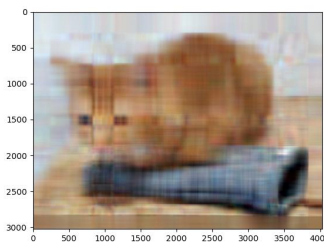
### 6.1 Low-Rank Approximation

The classical example of low-rank approximation is a image compression. More precisely, suppose we have an image like the below one:



**Figure 5:** The author's friend's cat Snyder.

It can be represented as three matrices  $A_R, A_G, A_B \in \mathbb{R}^{4032 \times 3024}$  corresponding to R, G, and B of the image. We perform a rank- $\ell$  approximation  $A_R = U_{R;\ell} \Sigma_{R;\ell} V_{R;\ell}^\top$ ,  $A_G = U_{G;\ell} \Sigma_{G;\ell} V_{G;\ell}^\top$ ,  $A_B = U_{B;\ell} \Sigma_{B;\ell} V_{B;\ell}^\top$  and then compose an image out of them, for different values of  $\ell$ . The results are shown below.

**(a)**  $\ell = 1$ .**(b)**  $\ell = 5$ .**(c)**  $\ell = 10$ .**(d)**  $\ell = 25$ .**(e)**  $\ell = 50$ .**(f)**  $\ell = 100$ .

By rank 100 approximation, the image is almost perfect. Now, the original image had  $3 \times 4032 \times 3024 = 36578304$  entries; at rank 100, we have  $3 \times 100 \times (4032 + 3024 + 1) = 2117100$  entries, so we need to store around 5% of the original image. Not bad!

See Appendix D for some code showing how these images were created.

## 6.2 PCA

Sadly, no cats for this example.

Suppose we, as course staff, have  $m$  students in our class, and  $n$  assignments. Let  $A \in \mathbb{R}^{m \times n}$  be a matrix, such that the  $i^{\text{th}}$  student's grade in the  $j^{\text{th}}$  assignment is  $A_{ij}$ .

- If we consider the assignments to be the data points, then  $A$  is a data matrix with column data.

Computing the eigenvectors  $\vec{u}_1, \dots, \vec{u}_m$  of  $AA^\top$  provides the principal components of the assignment data. We would expect something like:

- The first principal component  $\vec{u}_1$  probably indicates whether the assignment has lots of applications, or lots of theoretical problems. That is, the projection coefficient of an assignment  $\vec{x}$  onto the first principal component,  $\langle \vec{x}, \vec{u}_1 \rangle = (U^\top \vec{x})_1$ , would be a large positive number if the assignment  $\vec{x}$  is a purely theoretical assignment, a large negative number if the assignment  $\vec{x}$  is a purely application-based assignment, and somewhere in the middle if the assignment  $\vec{x}$  is somewhere in the middle.
  - The second principal component  $\vec{u}_2$  probably indicates whether the assignment is long or short, in the same way (looking at the value of the projection coefficient  $\langle \vec{x}, \vec{u}_2 \rangle = (U^\top \vec{x})_2$ ).
  - After a few more principal components, the data would be almost entirely contained in the principal component subspace, with the residuals being due to student performance on assignments not being perfectly consistent.
- If we consider the students to be the data points, then  $A$  is a data matrix with row data. Then the matrix  $A^\top$  is a data matrix with column data. Computing the eigenvectors  $\vec{v}_1, \dots, \vec{v}_n$  of  $(A^\top)(A^\top)^\top = A^\top A$  provides the principal components of the student data. We would expect something like:
    - The first principal component  $\vec{v}_1$  probably indicates whether the student prefers lots of applications, or lots of theoretical ideas. That is, the projection coefficient of a student  $\vec{y}$  onto the first principal component,  $\langle \vec{y}, \vec{v}_1 \rangle = (V^\top \vec{y})_1$ , would be a large positive number if the student  $\vec{y}$  purely favors theory, a large negative number if the student  $\vec{y}$  purely favors applications, and somewhere in the middle if the student  $\vec{y}$  is somewhere in the middle.
    - The second principal component  $\vec{v}_2$  probably indicates whether the student is an underclassman or an upperclassman, in the same way (looking at the value of the projection coefficient  $\langle \vec{y}, \vec{v}_2 \rangle = (V^\top \vec{y})_2$ ).
    - After a few more principal components, the data would be almost entirely contained in the principal component subspace, with the residuals being due to student performance on assignments not being perfectly consistent.

Given an assignment  $\vec{x} \in \mathbb{R}^m$  that we wanted to see "ground truth" student scores on, we would multiply it by  $U_{\ell_1} U_{\ell_1}^\top$  to get  $U_{\ell_1} U_{\ell_1}^\top \vec{x}$ , which is the "denoised" version of  $\vec{x}$ . Here the value of  $\ell_1$  can be picked by our validation scheme.

Similarly, given a student  $\vec{y} \in \mathbb{R}^n$  that we wanted to see "ground truth" assignment scores on, we would multiply it by  $V_{\ell_2} V_{\ell_2}^\top$  to get  $V_{\ell_2} V_{\ell_2}^\top \vec{y}$ , which is the "denoised" version of  $\vec{y}$ . Again, the value of  $\ell_2$  can be picked by our validation scheme. Note that  $\ell_1$  and  $\ell_2$  need not be equal.

*NOTE:* None of this section was based off of actual student data analysis; it was just the author's idea of what would be large orthogonal directions of variance in this dataset.

## 7 Final Comments

Low-rank approximation and principal component analysis are both powerful tools for data compression and data analysis. They both help us distill data down to its essential linear character, and so make it easy

to work with in future applications such as statistics and machine learning. PCA, in particular, is used as a powerful pre-processing step in data science.

## A Proof of Theorem 3

*Proof of Theorem 3.* The proof proceeds in several steps. Let  $A = U\Sigma V^\top$  throughout, and let  $\text{rank}(A) = r$ . Recall that we want to show that

$$A_k \in \underset{\substack{B \in \mathbb{R}^{m \times n} \\ \text{rank}(B) = \ell}}{\text{argmin}} \|A - B\|_F^2. \quad (21)$$

1. Reduce the problem with  $A$  down to a problem with  $\Sigma$ .

We use the fact that multiplication by an orthonormal matrix does not change the Frobenius norm (see [Note 11](#)) to get

$$\|A - B\|_F^2 = \|U\Sigma V^\top - B\|_F^2 \quad (22)$$

$$= \|U^\top(U\Sigma V^\top - B)V\|_F^2 \quad (23)$$

$$= \|U^\top U\Sigma V^\top V - U^\top B V\|_F^2 \quad (24)$$

$$= \|\Sigma - U^\top B V\|_F^2. \quad (25)$$

Now since  $U$  and  $V$  are orthonormal matrices, they are full rank, and so  $\text{rank}(B) = \text{rank}(U^\top B V)$ . Thus if we let  $X = U^\top B V$ , the problem reduces to showing that

$$\begin{bmatrix} \Sigma_k & 0_{k \times (n-k)} \\ 0_{(m-k) \times k} & 0_{(m-k) \times (n-k)} \end{bmatrix} \in \underset{\substack{X \in \mathbb{R}^{m \times n} \\ \text{rank}(X) = \ell}}{\text{argmin}} \|\Sigma - X\|_F^2. \quad (26)$$

2. Reduce the problem with  $\Sigma$  to a problem of finding the best matrix  $Q$  with orthonormal columns.

Write, where  $\vec{e}_i$  is the  $i^{\text{th}}$  standard basis vector in  $\mathbb{R}^m$ ,

$$\|\Sigma - X\|_F^2 = \sum_{i=1}^n \|(\Sigma - X)_i\|^2 \quad (27)$$

$$= \sum_{i=1}^n \|\sigma_i \vec{e}_i - \vec{x}_i\|^2. \quad (28)$$

Since  $\text{rank}(X) = \ell$ , the columns of  $X$  lie on an  $\ell$ -dimensional subspace, say  $S$  which is spanned by the orthonormal basis  $Q \in \mathbb{R}^{m \times \ell}$ . Within  $S$ , the closest  $\vec{x}_i$  to  $\sigma_i \vec{e}_i$  (i.e., the  $\vec{x}_i$  which minimizes  $\|\sigma_i \vec{e}_i - \vec{x}_i\|^2$ ) is the projection of  $\sigma_i \vec{e}_i$  onto  $S = \text{Col}(Q)$ , which is given by

$$\vec{x}_i = \text{proj}_S(\sigma_i \vec{e}_i) = \text{proj}_{\text{Col}(Q)}(\sigma_i \vec{e}_i) = \sigma_i Q Q^\top \vec{e}_i. \quad (29)$$

Simplifying the terms in the summand of the objective, we have

$$\|\sigma_i \vec{e}_i - \vec{x}_i\|^2 = \|\sigma_i \vec{e}_i - \sigma_i Q Q^\top \vec{e}_i\|^2 \quad (30)$$

$$= \sigma_i^2 \|\vec{e}_i - Q Q^\top \vec{e}_i\|^2. \quad (31)$$

Notice that, by the orthogonality principle of [Note 11](#), we have

$$0 = \langle \text{proj}_S(\vec{e}_i), \vec{e}_i - \text{proj}_S(\vec{e}_i) \rangle = \langle \text{proj}_{\text{Col}(Q)}(\vec{e}_i), \vec{e}_i - \text{proj}_{\text{Col}(Q)}(\vec{e}_i) \rangle = \langle Q Q^\top \vec{e}_i, \vec{e}_i - Q Q^\top \vec{e}_i \rangle. \quad (32)$$

Thus, expanding the squared norm,

$$\|\vec{e}_i - QQ^\top \vec{e}_i\|^2 = \langle \vec{e}_i - QQ^\top \vec{e}_i, \vec{e}_i - QQ^\top \vec{e}_i \rangle \quad (33)$$

$$= \langle \vec{e}_i, \vec{e}_i - QQ^\top \vec{e}_i \rangle - \underbrace{\langle QQ^\top \vec{e}_i, \vec{e}_i - QQ^\top \vec{e}_i \rangle}_{=0} \quad (34)$$

$$= \langle \vec{e}_i, \vec{e}_i - QQ^\top \vec{e}_i \rangle \quad (35)$$

$$= \langle \vec{e}_i, \vec{e}_i \rangle - \langle \vec{e}_i, QQ^\top \vec{e}_i \rangle \quad (36)$$

$$= \langle \vec{e}_i, \vec{e}_i \rangle - \langle Q^\top \vec{e}_i, Q^\top \vec{e}_i \rangle \quad (37)$$

$$= \|\vec{e}_i\|^2 - \|Q^\top \vec{e}_i\|^2 \quad (38)$$

$$= 1 - \|Q^\top \vec{e}_i\|^2. \quad (39)$$

Therefore, simplifying the original problem, we have

$$\operatorname{argmin}_{\substack{Q \in \mathbb{R}^{m \times \ell} \\ Q^\top Q = I_\ell}} \sum_{i=1}^n \sigma_i^2 \left( \|\vec{e}_i - QQ^\top \vec{e}_i\|^2 \right) = \operatorname{argmin}_{\substack{Q \in \mathbb{R}^{m \times \ell} \\ Q^\top Q = I_\ell}} \sum_{i=1}^n \sigma_i^2 \left( 1 - \|Q^\top \vec{e}_i\|^2 \right) \quad (40)$$

$$= \operatorname{argmax}_{\substack{Q \in \mathbb{R}^{m \times \ell} \\ Q^\top Q = I_\ell}} \sum_{i=1}^n \sigma_i^2 \|Q^\top \vec{e}_i\|^2 \quad (41)$$

$$= \operatorname{argmax}_{\substack{Q \in \mathbb{R}^{m \times \ell} \\ Q^\top Q = I_\ell}} \sum_{i=1}^n \|Q^\top (\sigma_i \vec{e}_i)\|^2 \quad (42)$$

$$= \operatorname{argmax}_{\substack{Q \in \mathbb{R}^{m \times \ell} \\ Q^\top Q = I_\ell}} \|Q^\top \Sigma\|_F^2. \quad (43)$$

since the squared Frobenius norm is the sum of the squared norms of the columns.

Thus, the following are equivalent problems:

- Find a rank- $\ell$  matrix  $X$  which minimizes  $\|\Sigma - X\|_F^2$ .
- Find a dimension- $\ell$  subspace  $S$  which minimizes  $\sum_{i=1}^n \|\sigma_i \vec{e}_i - \operatorname{proj}_S(\sigma_i \vec{e}_i)\|$ ;
- Find a matrix  $Q \in \mathbb{R}^{m \times \ell}$  such that  $Q^\top Q = I_\ell$ , which maximizes  $\|Q^\top \Sigma\|_F^2$ .

And so the third problem is the one we would like to solve.

- Reduce the problem of finding the best matrix  $Q$  with orthonormal columns to a problem with purely real numbers.

Write the columns of  $Q$  as  $Q = [\vec{q}_1 \ \cdots \ \vec{q}_\ell]$ . Then we can simplify the Frobenius norm as

$$\|Q^\top \Sigma\|_F^2 = \left\| \begin{bmatrix} \vec{q}_1^\top \\ \vdots \\ \vec{q}_\ell^\top \end{bmatrix} \begin{bmatrix} \sigma_1 \vec{e}_1 & \cdots & \sigma_n \vec{e}_n \end{bmatrix} \right\|_F^2 \quad (44)$$

$$= \left\| \begin{bmatrix} \sigma_1 \langle \vec{e}_1, \vec{q}_1 \rangle & \cdots & \sigma_n \langle \vec{e}_n, \vec{q}_1 \rangle \\ \vdots & \ddots & \vdots \\ \sigma_1 \langle \vec{e}_1, \vec{q}_\ell \rangle & \cdots & \sigma_n \langle \vec{e}_n, \vec{q}_\ell \rangle \end{bmatrix} \right\|_F^2 \quad (45)$$

$$= \sum_{i=1}^n \sum_{j=1}^{\ell} \sigma_i^2 \langle \vec{e}_i, \vec{q}_j \rangle^2 \quad (46)$$

$$= \sum_{i=1}^n \sum_{j=1}^{\ell} \sigma_i^2 Q_{ij}^2 \quad (47)$$

$$= \sum_{i=1}^n \sigma_i^2 \sum_{j=1}^{\ell} Q_{ij}^2 \quad (48)$$

$$= \sum_{i=1}^n \sigma_i^2 d_i \quad (49)$$

$$= \sum_{i=1}^r \sigma_i^2 d_i \quad (50)$$

where  $d_i := \sum_{j=1}^{\ell} Q_{ij}^2$ . The last simplification is because only the first  $r$  singular values are nonzero.

As the sum of squared numbers,  $d_i \geq 0$ . We show that  $d_i \leq 1$ . Indeed, let  $\widehat{Q}$  be the extension of  $Q$  to an orthonormal basis of  $\mathbb{R}^m$ ; in particular, define  $\widehat{Q} = \begin{bmatrix} Q & \widetilde{Q} \end{bmatrix}$  where  $\widetilde{Q} \in \mathbb{R}^{m \times (m-\ell)}$  has orthonormal columns, so that  $\widehat{Q} \in \mathbb{R}^{m \times m}$  is an orthonormal square matrix. Then  $\widehat{Q}$  has orthonormal rows (and columns), so each of its rows is unit norm, so

$$1 = \sum_{j=1}^m \widehat{Q}_{ij}^2 \quad (51)$$

$$= \sum_{j=1}^{\ell} \widehat{Q}_{ij}^2 + \sum_{j=\ell+1}^m \widehat{Q}_{ij}^2 \quad (52)$$

$$= \sum_{j=1}^{\ell} Q_{ij}^2 + \sum_{j=1}^{m-\ell} \widetilde{Q}_{ij}^2 \quad (53)$$

$$= d_i + \underbrace{\sum_{j=1}^{m-\ell} \widetilde{Q}_{ij}^2}_{\geq 0} \quad (54)$$

$$\implies d_i \leq 1. \quad (55)$$

Now, since  $Q$  is orthonormal, this also gives us a constraint on the  $d_i$ ; indeed,

$$\ell = \sum_{j=1}^{\ell} \|\vec{q}_j\|^2 \quad (56)$$

$$= \sum_{j=1}^{\ell} \sum_{i=1}^m Q_{ij}^2 \quad (57)$$

$$= \sum_{i=1}^m \sum_{j=1}^{\ell} Q_{ij}^2 \quad (58)$$

$$= \sum_{i=1}^m d_i. \quad (59)$$

This gives the problem:

$$\max_{d_1, \dots, d_m \in \mathbb{R}} \sum_{i=1}^r \sigma_i^2 d_i \quad (60)$$

$$\text{s.t. } d_i \geq 0 \quad i \in \{1, \dots, m\} \quad (61)$$

$$d_i \leq 1 \quad i \in \{1, \dots, m\} \quad (62)$$

$$\sum_{i=1}^m d_i = \ell. \quad (63)$$

Once we solve this problem, we can convert its answer onto constraints on  $Q$ . Any solution to this problem which has a corresponding  $Q$  will surely maximize  $\|\Sigma^\top Q\|_F^2$ ; it is left to solve this problem and prove that a solution has a corresponding  $Q$ .

4. Solve the numerical problem and find a maximizer.

One can show by a so-called **exchange argument**, or by inspection, that one maximizer of this problem is  $d_1^* = \dots = d_\ell^* = 1$  and  $d_{\ell+1}^* = \dots = d_m^* = 0$ , at which point the optimal value is

$$\sum_{i=1}^r \sigma_i^2 d_i^* = \sum_{i=1}^{\ell} \sigma_i^2. \quad (64)$$

5. Using the solution to the numerical optimization problem, find a  $Q_\star$  which maximizes  $\|Q^\top \Sigma\|_F^2$ .

Note that the quantity  $d_i$  is the squared norm of the  $i^{\text{th}}$  row of  $Q$ . Thus we are looking for a matrix  $Q_\star \in \mathbb{R}^{m \times \ell}$  which has:

- orthonormal columns;
- the 1<sup>st</sup> through  $\ell^{\text{th}}$  rows have unit norm;
- the  $(\ell + 1)^{\text{th}}$  through  $m^{\text{th}}$  rows are  $\vec{0}^\top$ .

A  $Q_\star$  which satisfies this is given by  $Q_\star = \begin{bmatrix} I_\ell \\ \mathbf{0}_{(m-\ell) \times \ell} \end{bmatrix}$ . Thus this  $Q_\star$  maximizes  $\|Q^\top \Sigma\|_F^2$  among all  $Q \in \mathbb{R}^{m \times \ell}$  with orthonormal columns.

6. Compute the original objective function using this  $Q_\star$  and show that it reduces to  $\|A - A_k\|_F^2$ .

We have that

$$\min_{\substack{B \in \mathbb{R}^{m \times n} \\ \text{rank}(B) = \ell}} \|A - B\|_F^2 = \min_{\substack{X \in \mathbb{R}^{m \times n} \\ \text{rank}(X) = \ell}} \|\Sigma - X\|_F^2 \quad (65)$$

$$= \|\Sigma - Q_\star Q_\star^\top \Sigma\|_F^2 \quad (66)$$

$$= \left\| \begin{bmatrix} \Sigma_r & \mathbf{0}_{r \times (n-r)} \\ \mathbf{0}_{(m-r) \times r} & \mathbf{0}_{(m-r) \times (n-r)} \end{bmatrix} - \begin{bmatrix} I_\ell \\ \mathbf{0}_{(m-\ell) \times \ell} \end{bmatrix} \begin{bmatrix} I_\ell \\ \mathbf{0}_{(m-\ell) \times \ell} \end{bmatrix}^\top \begin{bmatrix} \Sigma_r & \mathbf{0}_{r \times (n-r)} \\ \mathbf{0}_{(m-r) \times r} & \mathbf{0}_{(m-r) \times (n-r)} \end{bmatrix} \right\|_F^2 \quad (67)$$

$$= \left\| \begin{bmatrix} \Sigma_r & \mathbf{0}_{r \times (n-r)} \\ \mathbf{0}_{(m-r) \times r} & \mathbf{0}_{(m-r) \times (n-r)} \end{bmatrix} - \begin{bmatrix} \Sigma_\ell & \mathbf{0}_{\ell \times (n-\ell)} \\ \mathbf{0}_{(m-\ell) \times \ell} & \mathbf{0}_{(m-\ell) \times (n-\ell)} \end{bmatrix} \right\|_F^2 \quad (68)$$

$$= \left\| U \left( \begin{bmatrix} \Sigma_r & \mathbf{0}_{r \times (n-r)} \\ \mathbf{0}_{(m-r) \times r} & \mathbf{0}_{(m-r) \times (n-r)} \end{bmatrix} - \begin{bmatrix} \Sigma_\ell & \mathbf{0}_{\ell \times (n-\ell)} \\ \mathbf{0}_{(m-\ell) \times \ell} & \mathbf{0}_{(m-\ell) \times (n-\ell)} \end{bmatrix} \right) V^\top \right\|_F^2 \quad (69)$$

$$= \left\| U \Sigma V^\top - U \begin{bmatrix} \Sigma_\ell & \mathbf{0}_{\ell \times (n-\ell)} \\ \mathbf{0}_{(m-\ell) \times \ell} & \mathbf{0}_{(m-\ell) \times (n-\ell)} \end{bmatrix} V^\top \right\|_F^2 \quad (70)$$

$$= \|A - A_\ell\|_F^2. \quad (71)$$

Thus

$$A_\ell \in \underset{\substack{B \in \mathbb{R}^{m \times n} \\ \text{rank}(B) = \ell}}{\text{argmin}} \|A - B\|_F^2 \quad (72)$$

and the proof is complete.  $\square$

## B Proof of Theorem 4

*Proof of Theorem 4.* We first simplify each term in the objective function.

$$\|\vec{x}_i - WW^\top \vec{x}_i\|^2 = \langle \vec{x}_i - WW^\top \vec{x}_i, \vec{x}_i - WW^\top \vec{x}_i \rangle \quad (73)$$

$$= \langle \vec{x}_i, \vec{x}_i \rangle - \langle WW^\top \vec{x}_i, \vec{x}_i \rangle - \langle \vec{x}_i, WW^\top \vec{x}_i \rangle + \langle WW^\top \vec{x}_i, WW^\top \vec{x}_i \rangle \quad (74)$$

$$= \langle \vec{x}_i, \vec{x}_i \rangle - 2 \langle WW^\top \vec{x}_i, \vec{x}_i \rangle + \langle WW^\top WW^\top \vec{x}_i, \vec{x}_i \rangle \quad (75)$$

$$= \langle \vec{x}_i, \vec{x}_i \rangle - 2 \langle WW^\top \vec{x}_i, \vec{x}_i \rangle + \langle WW^\top \vec{x}_i, \vec{x}_i \rangle \quad (76)$$

$$= \langle \vec{x}_i, \vec{x}_i \rangle - \langle WW^\top \vec{x}_i, \vec{x}_i \rangle \quad (77)$$

$$= \langle \vec{x}_i, \vec{x}_i \rangle - \langle W^\top \vec{x}_i, W^\top \vec{x}_i \rangle \quad (78)$$

$$= \|\vec{x}_i\|^2 - \|W^\top \vec{x}_i\|^2. \quad (79)$$

Thus we can reduce the optimization of the objective to the simpler optimization

$$\underset{\substack{W \in \mathbb{R}^{n \times \ell} \\ W^\top W = I_\ell}}{\text{argmin}} \sum_{i=1}^n \|\vec{x}_i - WW^\top \vec{x}_i\|^2 = \underset{\substack{W \in \mathbb{R}^{n \times \ell} \\ W^\top W = I_\ell}}{\text{argmin}} \sum_{i=1}^n \left[ \|\vec{x}_i\|^2 - \|W^\top \vec{x}_i\|^2 \right] \quad (80)$$

$$= \underset{\substack{W \in \mathbb{R}^{n \times \ell} \\ W^\top W = I_\ell}}{\text{argmax}} \sum_{i=1}^n \|W^\top \vec{x}_i\|^2. \quad (81)$$

Now expanding the squared norm in terms of the sum of squares of each entry, we have

$$\sum_{i=1}^n \|W^\top \vec{x}_i\|^2 = \sum_{i=1}^n \sum_{k=1}^{\ell} (\vec{w}_k^\top \vec{x}_i)^2 \quad (82)$$

$$= \sum_{i=1}^n \sum_{k=1}^{\ell} \vec{w}_k^\top \vec{x}_i \vec{x}_i^\top \vec{w}_k \quad (83)$$

$$= \sum_{k=1}^{\ell} \sum_{i=1}^n \vec{w}_k^\top \vec{x}_i \vec{x}_i^\top \vec{w}_k \quad (84)$$

$$= \sum_{k=1}^{\ell} \vec{w}_k^\top \left( \sum_{i=1}^n \vec{x}_i \vec{x}_i^\top \right) \vec{w}_k \quad (85)$$



$$= \sum_{k=1}^{\ell} \vec{w}_k^{\top} (AA^{\top}) \vec{w}_k. \quad (86)$$

Computing  $AA^{\top} = U\Sigma\Sigma^{\top}U^{\top}$ , we have

$$\sum_{i=1}^n \left\| W^{\top} \vec{x}_i \right\|^2 = \sum_{k=1}^{\ell} \vec{w}_k^{\top} (AA^{\top}) \vec{w}_k \quad (87)$$

$$= \sum_{k=1}^{\ell} \vec{w}_k^{\top} U\Sigma\Sigma^{\top}U^{\top} \vec{w}_k. \quad (88)$$

We introduce the change of coordinates  $\vec{\tilde{w}}_k := U^{\top} \vec{w}_k$ . Note that since the  $\vec{w}_k$  are orthonormal, so too are the  $\vec{\tilde{w}}_k$ . Expanding out  $\Sigma\Sigma^{\top}$ , we have

$$\sum_{i=1}^n \left\| W^{\top} \vec{x}_i \right\|^2 = \sum_{k=1}^{\ell} \vec{\tilde{w}}_k^{\top} U\Sigma\Sigma^{\top}U^{\top} \vec{\tilde{w}}_k \quad (89)$$

$$= \sum_{k=1}^{\ell} \vec{\tilde{w}}_k^{\top} \Sigma\Sigma^{\top} \vec{\tilde{w}}_k \quad (90)$$

$$= \sum_{k=1}^{\ell} \vec{\tilde{w}}_k^{\top} \begin{bmatrix} \sigma_1^2 & & & & & & \\ & \ddots & & & & & \\ & & \sigma_r^2 & & & & \\ & & & 0 & & & \\ & & & & \ddots & & \\ & & & & & \ddots & \\ & & & & & & 0 \end{bmatrix} \vec{\tilde{w}}_k \quad (91)$$

(92)

Since the  $\vec{\tilde{w}}_k$ 's are orthonormal, a maximizing choice of  $\vec{\tilde{w}}_1, \dots, \vec{\tilde{w}}_{\ell}$  is  $\vec{e}_1, \dots, \vec{e}_{\ell}$  (i.e., the standard basis vectors)<sup>3</sup> – at which point the objective value is  $\sum_{k=1}^{\ell} \sigma_k^2$ . The corresponding  $\vec{w}_k$  is given by

$$\vec{w}_k = U\vec{\tilde{w}}_k = U\vec{e}_k = \vec{u}_k \quad (93)$$

so a maximizing  $W$  is  $U_{\ell}$  as desired.  $\square$

<sup>3</sup>There are other maximizers; for example, any permutation of the first  $\ell$  standard basis vectors, or if some singular values are equal then those can also be hit by an  $\vec{e}_k$ .

## C Code for PCA Plots

```
import numpy as np
import matplotlib.pyplot as plt
import pathlib

rng = np.random.RandomState(16)

n = 50
sigma = 0.1

s_vec = rng.randn(2, 1) # (2, 1)
X = s_vec @ rng.randn(1, n) + sigma * rng.randn(2, n) # (2, 50)
plt.axline((0, 0), s_vec.reshape(2, ), color='green')
plt.scatter(X[0], X[1])

plt.savefig(pathlib.Path("figures") / "pca_data.png")

U, S, Vh = np.linalg.svd(X)
plt.axline((0, 0), U[:, 0], color='orange')

plt.savefig(pathlib.Path("figures") / "pca.png")
```

## D Code for Low-Rank Compression

```
import numpy as np
import matplotlib.pyplot as plt
from matplotlib.image import imread
import pathlib

img = imread(pathlib.Path("figures") / "snyder.jpg")
A_R, A_G, A_B = img[:, :, 0], img[:, :, 1], img[:, :, 2]

# normalize to [0, 1]
A_R = A_R.astype(np.single) / 255
A_G = A_G.astype(np.single) / 255
A_B = A_B.astype(np.single) / 255

# Now take SVD and truncate to get approximations
U_R, S_R, Vh_R = np.linalg.svd(A_R)
U_G, S_G, Vh_G = np.linalg.svd(A_G)
U_B, S_B, Vh_B = np.linalg.svd(A_B)
```

```
Sigma_R, Sigma_G, Sigma_B = np.diag(S_R), np.diag(S_G), np.diag(S_B)
for l in (1, 5, 10, 25, 50, 100):
    img_l = np.zeros(shape=img.shape)
    img_l[:, :, 0] = U_R[:, :l] @ Sigma_R[:l, :l] @ Vh_R[:l]
    img_l[:, :, 1] = U_G[:, :l] @ Sigma_G[:l, :l] @ Vh_G[:l]
    img_l[:, :, 2] = U_B[:, :l] @ Sigma_B[:l, :l] @ Vh_B[:l]
    plt.imshow(img_l)
    plt.savefig(pathlib.Path("figures") / f"snyder_{l}.jpg")
```

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